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Boston College Working Papers in Economics, 2017

Originally posted on: <http://ideas.repec.org/p/boc/bocoec/937.html>

The Demand for Divisia Money: Theory and Evidence

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November 2017

Abstract: A money-in-the-utility function model is extended to capture the distinct roles of noninterest-earning currency and interest-earning deposits in providing liquidity services to households. It implies the existence of a stable money demand relationship that links a Divisia monetary aggregate to spending or income as a scale variable and the associated Divisia user-cost dual as an opportunity cost measure. Cointegrating money demand equations of this form appear in quarterly United States data spanning the period from 1967:1 through 2017:2, especially for the Divisia M2 aggregate. The identification of a stable money demand function over a period that includes the financial innovations of the 1980s and continues through the recent financial crisis and Great Recession suggests that a properly measured aggregate quantity of money can play a role in the conduct of monetary policy. That role can be of greater prominence when traditional interest rate policies are constrained by the zero lower bound.

JEL Codes: C43, E41.

Keywords: Divisia monetary aggregates, money demand, money-in-the-utility function.

Acknowledgments: Neither of us received any external support for, or has any financial interest that relates to, the research described in this paper.

Introduction

Since Milton Friedman's (1956) "restatement," the existence of a stable money demand function has been regarded as a necessary pre-condition for the success of any quantity-theoretic approach to monetary policy that would use information in broad monetary aggregates to achieve goals for aggregate spending or the price level. The general message of Friedman's essay, supported by the empirical papers that accompanied its publication, motivated a broad and active line of research in monetary economics that lasted more than three decades.¹ The primary focus of this agenda was the search for statistical relationships that linked the demand for real money balances to a small number of determinants, especially income or spending and interest rates. By the end of the 1970s, considerable evidence had been assembled to support the notion of a stable demand for money function and, when inflation had risen to increasingly higher rates, this evidence provided the foundation for a monetary policy strategy based on intermediate targets for growth rates of the money supply.

Interest in the demand for money, however, quickly waned in the early 1990s, beginning with the publication of highly influential studies by Bernanke and Blinder (1992) and Friedman and Kuttner (1992). Those studies showed that previously stable relationships between money and other key macroeconomic variables – including popular money demand specifications – weakened considerably after 1980. Moreover, their results indicated that any significant effects of monetary policy actions on aggregate activity were manifested through changes in the federal funds rate rather than the quantity of money.² Conventional wisdom held that sweeping financial innovations during the 1980s permanently altered whatever associations might have been found in the data of the 1960s and 70s. Although the findings in these studies and others like them subsequently were shown by Belongia (1996) and, more recently, Hendrickson (2014), to depend partly on their use of the Federal Reserve's simple-sum monetary aggregates instead of the Divisia alternatives proposed by Barnett (1980), they nevertheless had led a professional consensus to coalesce around the idea that measures of money could be safely excluded from the information set used to evaluate the effects of monetary policy on economic activity or to assess the relative ease or tightness of policy at a moment in time.

About the same time, John Taylor also shifted his focus from money to interest rates. Having earlier (Taylor 1979) investigated the properties of a monetary policy rule based on

¹ Friedman's essay and the accompanying papers appeared in *Studies in the Quantity Theory of Money* (1956). Laidler (1993) provides a survey of the research on the demand for money that followed.

² A related issue, also associated with measurement of the money supply, was the so-called "velocity problem." For example, the sharp decline in the velocity of M2 (a shift in the demand for M2) after the financial innovations of the 1980s quickly undercut the perceived usefulness of the P-Star model of Hallman, Porter, and Small (1991).

money, he now (Taylor 1993) suggested that Federal Reserve policy, beginning in the late 1980s, could be adequately described by a strikingly simple alternative based on the federal funds rate. Notably, this rule called for adjustments in the federal funds rate in response to movements in inflation and output alone with no separate role for the money supply. Interest rate rules of the general form proposed by Taylor then were incorporated into the New Keynesian models developed by Clarida, Galí, and Gertler (1999) and Woodford (2003). And though Nelson (2008) and even Taylor (1996) himself showed that New Keynesian models with interest rate rules are not inconsistent with the quantity-theoretic view that inflation gets determined, in the long run, by the growth rate of the money supply, these models also served to illustrate how monetary policy analysis could be conducted within a theoretically coherent and internally consistent framework that makes no explicit reference to either the demand for or supply of money. As such, models of this type also excluded any independent transmission mechanism by which variations in the quantity of money can affect real activity independent of any influence associated with variations in interest rates.

Recent events, however, provide good reason to reconsider the role of the monetary aggregates in monetary policy analysis. Most significantly, the zero lower bound on nominal interest rates, which became a binding constraint on the Federal Reserve's interest rate policy over an extended period from 2008 through 2015, highlights the limitations of any approach that uses the federal funds rate alone to gauge the stance of monetary policy. If, for example, Federal Reserve actions affect spending and prices through changes in variables other than the funds rate, these effects will not be captured by models that exclude, by construction, any channel of monetary transmission apart from an interest rate. In fact, empirical evidence shows how Divisia monetary aggregates can be used within structural vector autoregressive time series models to help identify monetary policy shocks before and during the period of zero nominal interest rates by capturing the consequences of "unconventional" policy actions such as large-scale asset purchases or "quantitative easing" through their effects on the growth rate of Divisia money.³ Because the results from these studies are consistent with the ideas that helped motivate the money demand research agenda sixty years ago, these same Divisia monetary aggregates might be used to estimate stable money demand relationships of the kind envisioned by Friedman (1956). If identified, these relationships then could serve as the foundations for a quantity-theoretic approach to monetary policymaking and analysis that downplays the significance of the zero lower bound, yet also works reliably to stabilize inflation during normal periods with positive nominal interest rates.

To explore this possibility, this paper begins by modifying Lucas' (2000) version of the money-in-the-utility function model, developed originally by Sidrauski (1967) and Brock (1974),

³ See, for example, Keating, Kelly, Smith, and Valcarcel (2014) and Belongia and Ireland (2015b, 2016, 2017),

by introducing separate roles for noninterest-earning currency and interest-earning deposits in providing a representative household with liquidity services that allow it to purchase goods and services at the expense of less time and effort. This extension to the theory makes clear that the money demand relationship implied by the model applies to a Divisia monetary aggregate but not a simple-sum measure of the type provided officially by the Federal Reserve. The same theoretical extension also reinforces Belongia's (2006) argument that the price, or user-cost, dual to the Divisia monetary aggregate ought to appear in place of a short-term nominal interest rate as a preferred opportunity cost measure in the money demand equation. Finally, the extended money-in-the-utility function model can be used to motivate renewed interest in classic empirical specifications for money demand, originally proposed by Cagan (1956), Selden (1956), Latané (1960), and Meltzer (1963), adapted to apply to the Divisia monetary aggregates instead of the standard simple sum measures.

The paper goes on to estimate these money demand equations using Johansen's (1991) maximum likelihood approach together with quarterly data on the Divisia aggregates described by Barnett, Liu, Mattson, and van den Noort (2013) and made available through the Center for Financial Stability's (CFS) website. In both the full sample of data running from 1967:1 through 2017:2 and a subsample beginning in 1983:1, cointegrating money demand relationships of the form suggested by the theory link the Divisia monetary aggregates to either spending or income as a scale variable and the associated Divisia user cost aggregate as an opportunity cost measure. Estimates of the elasticity of Divisia money demand with respect to changes in the user cost decline when estimated with data from the post-1983 subsample, relative to the full sample. But the long-run money demand relationships appear most consistently for the Divisia M2 aggregate, implying that a monetary policy strategy focused on stabilizing Divisia M2 growth has the potential to stabilize the aggregate price level or its rate of change.⁴

By using the most recent data to estimate statistical money demand equations, this paper joins with a few others, including Judson, Schlusche, and Wong (2014), Lucas and Nicolini (2015), and Anderson, Bordo, and Duca (2017), that also are motivated by the financial crisis and period of zero nominal interest rates to reconsider a more traditional, quantity-theoretic approach to monetary policy analysis. Meanwhile, Benati, Lucas, Nicolini, and Weber (2016) and Benati (2017) revive the analysis of long-run money demand by estimating these equations using international data extending back, in some cases, to the 19th century. Likewise, Serletis and Gogas (2014) estimate long-run money demand relationships for the CFS Divisia aggregates, but use the three-month United States Treasury bill rate as their

⁴ Belongia and Ireland (2015a) present evidence on the viability of targeting a path for nominal GDP by controlling the behavior of a Divisia monetary aggregate, a result suggestive of a stable demand for money function.

opportunity cost measure. Because the theory suggests, however, that the “price” of monetary services is the user cost dual to the economic quantity aggregate, the empirical specifications that follow use this measure instead. The results presented here, therefore, provide evidence of stable money demand relationships based on price and quantity data derived from the same principles.

Theory

The Model

As noted above, the model developed here builds further on Lucas’ (2000) variant of the money-in-the-utility function models of Sidrauski (1967) and Brock (1974) by introducing separate roles for noninterest-earning currency and interest-earning deposits in providing a representative household with liquidity services that allow it to purchase goods and services with less effort. Through this extension, the theory makes clear that the aggregate of currency and deposits appearing in a properly-specified money demand relationship is a Divisia aggregate, and not simple sum measures like those constructed, officially, by the Federal Reserve. While the description of household optimization provided here could be incorporated into a dynamic, stochastic, general equilibrium model along the same lines followed by Belongia and Ireland (2014) and Ireland (2014), the perfect foresight, partial equilibrium framework used here, instead, simplifies the analysis by abstracting from unnecessary general equilibrium considerations and highlights, as well, that the basic properties of the money demand relationship derived here do not depend on the details of how other sectors of the economy might be modeled.

Also for simplicity, the theoretical analysis proceeds here under the assumption that a representative household substitutes between currency and a single type of interest-earning bank deposit in its efforts to construct the portfolio of these two liquid assets that most efficiently provides the monetary services it uses in making transactions during each period. Additional types of deposits, all paying interest at different rates and each playing its own role in the household’s portfolio of monetary assets, could easily be incorporated, at the cost of requiring slightly more detailed notation and more tedious algebra. The empirical work, by contrast, uses various Divisia monetary aggregates that do include a wide range of monetary assets available in the United States today.

In the model, an infinitely-lived representative household enters each period $t = 0, 1, 2, \dots$ with M_{t-1} units of currency and B_{t-1} bonds. At the very start of a beginning-of-period asset trading and allocation session, the household receives T_t additional units of currency. Next, the household’s bonds mature, providing B_{t-1} more units of currency. The

household uses some of its currency to purchase B_t new bonds at the price of $1/(1+r_t)$ dollars per bond, where r_t denotes the net nominal interest rate between t and $t+1$. The household divides its remaining currency into an amount N_t to be used for transactions purposes during the period and an amount to be deposited in an interest-earning bank account instead. At the same time, the household also borrows L_t dollars from the bank, which then gets credited to its bank account. Thus, the total value D_t of the household's deposits must satisfy

$$M_{t-1} + T_t + B_{t-1} - \frac{B_t}{1+r_t} - N_t + L_t \geq D_t \quad (1)$$

for all $t = 0, 1, 2, \dots$.

To describe its activities following this initial asset allocation session, it is helpful to follow Lucas (1980) by imagining that the representative household consists of two members: a worker and a shopper. During each period $t = 0, 1, 2, \dots$, the representative household's worker supplies labor inelastically to produce y_t units of output, where y_t may fluctuate or grow over time, reflecting any arbitrary pattern of exogenous technological change, so long as it is anticipated in line with the perfect foresight assumption. The worker then sells this output to shoppers from other households at the nominal price of P_t dollars per unit of the good. The representative household's shopper, meanwhile, purchases c_t units of the good at the same price P_t from workers of other households.

At the end of each period $t = 0, 1, 2, \dots$, the household's two members reunite to consume the shopper's purchases. The household's preferences over the infinite horizon are described by the utility function

$$\sum_{t=0}^{\infty} \beta^t \left(\frac{1}{1-\sigma} \right) \left\{ \left[c_t v \left(\frac{M_t^a}{P_t c_t} \right) \right]^{1-\sigma} - 1 \right\}, \quad (2)$$

where the discount factor lies between zero and one, so that $0 < \beta < 1$, and the intertemporal elasticity of substitution parameter is positive, so that $\sigma > 0$. In the special case where $\sigma = 1$, the single-period utility function in (2) should be understood to take its limiting form

$$\ln(c_t) + \ln \left[v \left(\frac{M_t^a}{P_t c_t} \right) \right].$$

In (2), the function v captures the time-and-effort-saving services provided to the representative household's shopper by the monetary aggregate M_t^a , formed from currency and deposits according to

$$g(N_t, D_t) \geq M_t^a,$$

where the monetary aggregator g is assumed to be homogeneous of degree one, so that the underlying transaction technology exhibits constant returns to scale and so that this constraint can be rewritten equivalently in real terms as

$$g\left(\frac{N_t}{P_t}, \frac{D_t}{P_t}\right) \geq \frac{M_t^a}{P_t} \quad (3)$$

for all $t = 0, 1, 2, \dots$. As noted by Lucas (2000), the specific form of the utility function in (2) makes the model consistent with balanced growth: if the household's real income y_t grows at a constant long-run growth rate, then its optimal choices of consumption c_t and the real monetary aggregate M_t^a / P_t will grow at the same long-run rate.

The specification in (3) allows for a wide range of possibilities regarding the degree of substitutability between currency and deposits in providing the household with monetary services. Belongia and Ireland (2014) and Ireland (2014) assume, for example, that g takes the constant elasticity of substitution form

$$g(N_t, D_t) = [\nu^{1/\omega} N_t^{(\omega-1)/\omega} + (1-\nu)^{1/\omega} D_t^{(\omega-1)/\omega}]^{\omega/(\omega-1)},$$

with $\omega > 0$ and $0 < \nu < 1$. Here, this CES function could serve as a special case, but the analysis applies more generally to any aggregator g satisfying the homogeneity assumption that doubling the quantities of currency and deposits held also doubles the monetary services produced by the combination of these underlying assets. Note, however, that another special case,

$$g(N_t, D_t) = N_t + D_t,$$

also is subsumed under the more general formulation allowed for in (3) and is consistent with the official practice of constructing monetary aggregates as the "simple sums" of funds held as currency and deposits. It implies, however, that so long as deposits pay interest while currency does not, the household will choose a corner solution in which it holds only deposits and no currency. In economies, including the United States', where households are observed to hold both currency and deposits, this simple-sum aggregator appears far less plausible than others in which currency and deposits substitute imperfectly for one another in providing monetary services. This is the same point made originally by Barnett (1980): unless currency and deposits are treated, counterfactually, as perfect substitutes, simple-sum monetary

aggregation is inconsistent with economic theory, and simple-sum monetary aggregates are unlikely to measure accurately the true flows of monetary services demanded by households.

Also at the end of each period $t = 0, 1, 2, \dots$, the household receives an interest payment $r_t^d D_t$ on the deposits it holds during the period, but must also repay with interest $r_t^l L_t$ the loans it received earlier from the bank. After accounting for these receipts and payments, as well as the income $P_t y_t$ earned and funds $P_t c_t$ spent during the period, the household carries M_t units of currency into period $t + 1$, where

$$N_t + (1 + r_t^d)D_t + P_t y_t - P_t c_t - (1 + r_t^l)L_t \geq M_t \quad (4)$$

for all $t = 0, 1, 2, \dots$. The household, therefore, chooses B_t , N_t , D_t , L_t , c_t , M_t^a , and M_t for all $t = 0, 1, 2, \dots$ to maximize the utility function in (2) subject to the constraints in (1), (3), and (4), each of which must hold for all $t = 0, 1, 2, \dots$.

The most convenient way of characterizing the solution to the household's problem starts by rewriting the constraints from (1) and (4) in real terms, as

$$\frac{M_{t-1} + T_t + B_{t-1} + L_t}{P_t} \geq \frac{B_t / (1 + r_t) + N_t + D_t}{P_t} \quad (5)$$

and

$$\frac{N_t + (1 + r_t^d)D_t}{P_t} + y_t \geq c_t + \frac{(1 + r_t^l)L_t + M_t}{P_t}, \quad (6)$$

before introducing Λ_t^1 , Λ_t^2 , and Λ_t^3 as the nonnegative Lagrange multipliers on (3), (5), and (6).

The first-order conditions for the household's problem can then be written as

$$\frac{\Lambda_t^2}{P_t(1 + r_t)} = \frac{\beta \Lambda_{t+1}^2}{P_{t+1}}, \quad (7)$$

$$\Lambda_t^1 g_1 \left(\frac{N_t}{P_t}, \frac{D_t}{P_t} \right) = \Lambda_t^2 - \Lambda_t^3, \quad (8)$$

$$\Lambda_t^1 g_2 \left(\frac{N_t}{P_t}, \frac{D_t}{P_t} \right) = \Lambda_t^2 - (1 + r_t^d) \Lambda_t^3, \quad (9)$$

$$\Lambda_t^2 = (1 + r_t^l) \Lambda_t^3, \quad (10)$$

$$\left[c_t v \left(\frac{M_t^a}{P_t c_t} \right) \right]^{-\sigma} \left[v \left(\frac{M_t^a}{P_t c_t} \right) - \left(\frac{M_t^a}{P_t c_t} \right) v' \left(\frac{M_t^a}{P_t c_t} \right) \right] = \Lambda_t^3, \quad (11)$$

$$\left[c_t v \left(\frac{M_t^a}{P_t c_t} \right) \right]^{-\sigma} v' \left(\frac{M_t^a}{P_t c_t} \right) = \Lambda_t^1, \quad (12)$$

and

$$\frac{\Lambda_t^3}{P_t} = \frac{\beta \Lambda_{t+1}^2}{P_{t+1}}, \quad (13)$$

together with (3), (5), and (6) as equalities for all $t = 0, 1, 2, \dots$.

Monetary Aggregation and Money Demand

Note that (7) and (13) imply

$$\Lambda_t^2 = (1 + r_t) \Lambda_t^3. \quad (14)$$

Comparing (10) and (14) then reveals that the absence of arbitrage requires $r_t = r_t^l$ for all $t = 0, 1, 2, \dots$, so that the interest rate on bonds and loans are always equal. This result reflects the fact that in the model, for simplicity, bonds and loans are both risk-free assets that provide no monetary services. In considering the monetary aggregates implied by this model, therefore, either the bond rate or the loan rate serves equally well as the “benchmark” against which the lower interest rate on deposits and the zero interest rate on currency can be compared in computing their associated user costs. In the United States’ economy, however, short-term risk-free bonds – for example, three-month United States Treasury bills – may provide liquidity services and, indeed, are included in the broadest, Divisia M4 aggregate defined by Barnett, Liu, Mattson, and van den Noort (2013). By contrast, bank loans are not included in any of the CFS Divisia monetary aggregates and, according to Barnett, Liu, Mattson, and van den Noort (2013), an interest rate on commercial and industrial loans serves for much of the sample period as the benchmark rate in constructing these aggregates. Thus, while neither bonds nor bank loans play a direct role in this simple model in providing monetary services to households or borrowed funds to either the government or private firms, both are included in the household’s problem to point out that the model’s benchmark rate could be either the rate r_t on illiquid bonds or the rate r_t^l on illiquid bank loans.

Given r_t (or r_t^l), define r_t^a so that

$$r_t - r_t^a = \Lambda_t^1 / \Lambda_t^3 \quad (15)$$

for all $t = 0, 1, 2, \dots$. Using (14) and (15), it can be verified that if the choices of currency and deposits are not of independent interest, the household’s problem can be stated more simply as

one of choosing B_t , L_t , c_t , M_t^a , and M_t for all $t=0,1,2,\dots$ to maximize the utility function in (2) subject to the constraints

$$\frac{M_{t-1} + T_t + B_{t-1} + L_t}{P_t} \geq \frac{B_t / (1 + r_t) + M_t^a}{P_t}$$

and

$$\frac{(1 + r_t^a)M_t^a}{P_t} + y_t \geq c_t + \frac{(1 + r_t^l)L_t + M_t}{P_t}$$

for all $t=0,1,2,\dots$, confirming that M_t^a can be treated as a true microeconomic aggregate of monetary services, with “own rate of return” equal to r_t^a . Equation (3) still makes clear, however, that this true monetary aggregate will not correspond to a simple sum of currency and deposits, except in the counterfactual case where these assets are perfect substitutes.

Equations (14) and (15) also allow (8) and (9) to be rewritten as

$$g_1\left(\frac{N_t}{P_t}, \frac{D_t}{P_t}\right) = \frac{r_t \Lambda_t^3}{\Lambda_t^1} = \frac{r_t}{r_t - r_t^a} = \frac{u_t^n}{u_t^a} \quad (16)$$

and

$$g_2\left(\frac{N_t}{P_t}, \frac{D_t}{P_t}\right) = \frac{(r_t - r_t^d) \Lambda_t^3}{\Lambda_t^1} = \frac{r_t - r_t^d}{r_t - r_t^a} = \frac{u_t^d}{u_t^a}, \quad (17)$$

where the last equalities in each of these expressions use Barnett’s (1978) formula to define the user costs of currency and deposits as

$$u_t^n = \frac{r_t}{1 + r_t}$$

and

$$u_t^d = \frac{r_t - r_t^d}{1 + r_t}$$

and the user cost of the monetary aggregate M_t^a as

$$u_t^a = \frac{r_t - r_t^a}{1 + r_t}. \quad (18)$$

Using (16) and (17) together with Euler’s theorem, the homogeneity of the monetary aggregator g can now be seen to imply

$$\frac{M_t^a}{P_t} = g\left(\frac{N_t}{P_t}, \frac{D_t}{P_t}\right) = \left(\frac{N_t}{P_t}\right) g_1\left(\frac{N_t}{P_t}, \frac{D_t}{P_t}\right) + \left(\frac{D_t}{P_t}\right) g_2\left(\frac{N_t}{P_t}, \frac{D_t}{P_t}\right) = \frac{u_t^n}{u_t^a} \left(\frac{N_t}{P_t}\right) + \frac{u_t^d}{u_t^a} \left(\frac{D_t}{P_t}\right)$$

or, more simply,

$$u_t^a M_t^a = u_t^n N_t + u_t^d D_t. \quad (19)$$

The left-hand side of (19) measures total expenditures on monetary services; the right-hand side decomposes these expenditures into components provided by currency and deposits.

Although neither of the two terms on the left-hand side, the user cost u_t^a nor the quantity aggregate M_t^a , is observable individually, all of the terms on the right-hand side are observable from data on currency, deposits, and the interest rates on deposits and the benchmark, illiquid asset. Thus, total expenditures on the left-hand side can be inferred from the sum on the right, and the expenditure shares for currency

$$s_t^n = \frac{u_t^n N_t}{u_t^a M_t^a} = \frac{u_t^n N_t}{u_t^n N_t + u_t^d D_t}$$

and for deposits

$$s_t^d = \frac{u_t^d D_t}{u_t^a M_t^a} = \frac{u_t^d D_t}{u_t^n N_t + u_t^d D_t}$$

are observable as well.

In practice, therefore, the measurement problem posed by the theory involves distinguishing between the two terms u_t^a and M_t^a on the left-hand side of (19). This problem, however, is conceptually no different from the one that arises in a national income accounting exercise whereby nominal GDP is computed as the sum of all dollar spending on finished goods and services produced in the economy during a given period of time, but must then be broken down into its two components – real GDP and the aggregate price level – using economic aggregation and index-number theory. For monetary aggregation, in particular, Barnett (1980, 2012) shows that the discrete-time Divisia approximation

$$\ln(M_t^a) - \ln(M_{t-1}^a) \approx \left(\frac{s_t^n + s_{t-1}^n}{2} \right) [\ln(N_t) - \ln(N_{t-1})] + \left(\frac{s_t^d + s_{t-1}^d}{2} \right) [\ln(D_t) - \ln(D_{t-1})]$$

tracks the true quantity aggregate closely under a wide range of circumstances and, importantly, does not require knowledge of the form or parameters of the monetary aggregator function g . Moreover, once a series for M_t^a is constructed using this formula, a

corresponding series for the user cost u_t^a can be computed, as well, based on “factor reversal,” that is, by dividing total expenditures $u_t^a M_t^a$ by the quantity index M_t^a . This is, in fact, exactly how the CFS Divisia quantity and user costs are constructed for monetary aggregates that include currency and a wide range of bank deposits and other highly liquid assets.

Hence, this model provides a theory of the demand for money as measured by a Divisia aggregate but not a simple sum measure. To see this, combine (11), (12), and (15) to obtain

$$r_t - r_t^a = \frac{v' \left(\frac{M_t^a}{P_t c_t} \right)}{v \left(\frac{M_t^a}{P_t c_t} \right) - \left(\frac{M_t^a}{P_t c_t} \right) v' \left(\frac{M_t^a}{P_t c_t} \right)}. \quad (20)$$

This condition takes the same form as Lucas' (2000, p.256) equation 3.7, except that M_t^a in (20) is clearly best measured by a Divisia monetary aggregate and the additional, "own rate" term r_t^a appears on the left-hand side to account for the interest paid on many forms of bank deposits. The equation defines, implicitly, a money demand function of the form

$$\frac{M_t^a}{P_t c_t} = h(r_t - r_t^a) \quad (21)$$

with unitary scale elasticity and $r_t - r_t^a$ as its opportunity cost measure. Equation (18) links this opportunity cost to the user-cost dual u_t^a of the Divisia quantity aggregate. With regard to the empirical specifications that follow, Belongia (2006) notes that the Divisia user cost, as a theoretically-coherent measure of the own-price of monetary services in a money demand relationship, also implies that an interest rate, which is associated with the price of bonds, represents the price of a substitute, rather than the opportunity cost of holding money.⁵

Three Empirical Specifications

Let

$$m_t^a = \frac{M_t^a}{P_t c_t} \quad (22)$$

denote the ratio of the real monetary aggregate to consumption spending c_t , and invert (21) to obtain

$$r_t - r_t^a = \psi(m_t^a).$$

Then (20) implies that, given ψ , the original utility function v will solve the differential equation

⁵ Note that the explicit derivation of money's own price as the dual to the economic quantity aggregate stands in contrast to the empirical tradition of measuring the opportunity of holding money balances by the *ad hoc* choice of a short term interest rate, a long rate, an interest rate spread, or a wage rate.

$$\frac{v'(m^a)}{v(m^a)} = \frac{\psi(m^a)}{1 + m^a \psi(m^a)}, \quad (23)$$

which coincides with Lucas' (2000, p.257) equation 3.9 except that, again, the functions v and ψ have, as their arguments, the ratio of a real Divisia monetary aggregate to consumption.

Equations (21) and (23) can be used to specialize the model so that it motivates several classical empirical formulations for money demand, modified here to relate the demand for a Divisia monetary aggregate to its associated user cost. Suppose, for example, that the money demand function (21) takes the “double-log” form proposed by Meltzer (1963):

$$\ln(M_t^a / P_t) = \alpha_0 + \ln(c_t) - \alpha_1 \ln(r_t - r_t^a), \quad (24)$$

so that $\alpha_1 > 0$ measures the absolute value of the constant elasticity of real Divisia money demand with respect to changes in its opportunity cost $r_t - r_t^a$. Equation (23) then can be used to verify that this specification for money demand is implied by any utility function of the form

$$v(m^a) = \alpha \left[1 + \frac{\exp(\alpha_0 / \alpha_1)}{(m^a)^{(1-\alpha_1)/\alpha_1}} \right]^{-\alpha_1/(1-\alpha_1)},$$

with $\alpha > 0$. Alternatively, if (21) takes the “semi-log” form proposed by Cagan (1956),

$$\ln(M_t^a / P_t) = \delta_0 + \ln(c_t) - \delta_1 (r_t - r_t^a), \quad (25)$$

where $\delta_1 > 0$ measures the absolute value of the constant semi-elasticity of real Divisia money demand with respect to the opportunity cost, (23) requires that

$$\frac{v'(m^a)}{v(m^a)} = \frac{\delta_0 - \ln(m^a)}{\delta_1 + \delta_0 m^a - m^a \ln(m^a)}.$$

This differential equation lacks a closed-form solution, but for any specific values of δ_0 and δ_1 obtained by estimating (25), $v(m^a)$ can be characterized by solving it numerically.

Ireland (2009) estimates money demand functions of the double-log and semi-log forms (24) and (25) using data on simple-sum M1 and the three-month United States Treasury bill rate. More recently, Benati, Lucas, Nicolini, and Weber (2016) and Benati (2017) use long-run data on simple-sum M1, augmented for the period since 1980 by adding balances from money market deposit accounts, to estimate another classic money demand specification, first proposed by Selden (1956) and Latané (1960). When applied to simple-sum aggregates, this specification posits a linear relationship between monetary velocity and a short-term nominal interest rate. Here, the same specification is adapted for use with the Divisia monetary aggregates by letting

$$v_t^a = \frac{P_t c_t}{M_t^a} \quad (26)$$

denote the consumption velocity of M_t^a and assuming that (21) implies

$$v_t^a = \gamma_0 + \gamma_1(r_t - r_t^a), \quad (27)$$

where the parameter $\gamma_1 > 0$ measures the constant responsiveness of Divisia monetary velocity to changes in the opportunity cost measure. In this case, (23) has a solution of the form

$$v(m^a) = \gamma(m^a)^{1/(1+\gamma_1)}(\gamma_0 m^a - 1 - \gamma_1)^{\gamma_1/(1+\gamma_1)},$$

with $\gamma > 0$.

Thus, depending on the specific form of the utility function v , the model presented here is consistent with any of the classic money demand functions (24), (25), or (27). The next section presents estimates of these three money demand equations, based on United States data on Divisia monetary aggregates and their associated user costs.

Evidence

Overview of the Data

Figures 1 and 2 illustrate the behavior of the four variables appearing in the three Divisia money-demand specifications (24), (25) and (27): the natural logarithm of the ratio of a Divisia monetary aggregate M_t^a to nominal personal consumption expenditures $P_t c_t$, the natural logarithm and the level of the associated opportunity cost measure $r_t - r_t^a$, and the consumption velocity of money computed by dividing nominal personal consumption expenditures by the Divisia monetary aggregate. As noted previously, the series on Divisia money are described by Barnett, Liu, Mattson, and van den Noort (2013) and made available through the Center for Financial Stability's website. Because the theory links money demand most closely to consumption, personal consumption expenditures is used as the scale variable in figures 1 and 2 and in generating the benchmark statistical results that follow. Figures and tables in the appendix show, however, that very similar results and conclusions follow when nominal GDP is used to scale the data on Divisia money instead. Both series – for nominal PCE and GDP – are drawn from the Federal Reserve Bank of St. Louis' FRED database. The data on aggregate spending and income dictate a quarterly frequency for all series, while the starting point for the CFS monetary data determines the sample period, running from 1967:1 through 2017:2.

Four levels of monetary aggregation are considered. The Divisia M1 and M2 measures contain the same assets included in the Federal Reserve's official, simple-sum aggregates. In particular, Divisia M1 includes currency, travelers' checks, demand deposits, and other

checkable deposits. Divisia M2 adds savings deposits, including money market deposit accounts, retail money market mutual fund shares, and small time deposits. The MZM monetary aggregate, which excludes the small time deposit component of M2 but adds institutional money market mutual fund shares, was first proposed by Motley (1988) and given the label “money, zero maturity” by Poole (1991). The CFS data include a Divisia measure at this level of aggregation as well. Finally, Divisia M4 – the broadest aggregate compiled by the CFS – combines all of the assets in M2 and MZM with large time deposits, overnight and term repurchase agreements, commercial paper, and United States Treasury bills to obtain a collection similar to that included in the Federal Reserve’s discontinued L measure of liquidity.

The CFS also produces user cost aggregates for each of the Divisia quantity aggregates, constructed via factor reversal by dividing total expenditures on monetary services, measured according to an extended version of the formula displayed in (19), but with multiple classes of deposits (and in the case of M4, other highly liquid assets), by the corresponding quantity aggregate. Because the Divisia user cost of each aggregate is defined and measured according to the formula in (18), whereas the money demand relationship (21) implied by the theory links money demand, instead, to the opportunity cost measure $r_t - r_t^a$, each CFS user cost series is multiplied by $1 + r_t$, where r_t is the benchmark rate also described by Barnett, Liu, Mattson, and van den Noort (2013), before use in the statistical analysis. In all of the discussion that follows, therefore, this measure $r_t - r_t^a$ is referred to interchangeably as the “user cost” or “opportunity cost” of the corresponding Divisia monetary aggregate.

Figure 1 plots each quarterly series in isolation, to illustrate its univariate properties. The panels in the first column show that at all four levels of aggregation, the log Divisia money-consumption ratio trends steadily downward during the “Great Inflation” of the 1970s before leveling off and remaining stable from the early 1980s through the onset of the financial crisis in 2007-08. Since then, all four Divisia aggregates – but especially M1, M2, and MZM – have grown more rapidly than spending, generating a new, upward trend in the money-consumption ratio.

The next two columns of figure 1 plot the opportunity cost variable $r_t - r_t^a$ in logs, as it appears in Meltzer’s (1963) specification (24) and in levels, as it appears in Cagan’s (1956) model (25). Regardless of the transformation, these user-cost variables display upward and then downward trends that mirror those in the money-consumption ratio, as one would expect if long-run money demand was described by one of these relations. At the same time, however, the graphs make clear that the user cost series display much larger transitory deviations from their long-run trends.

The series for the Divisia consumption velocities displayed in the last column of figure 1 and that appear in the alternative money demand model (27) proposed by Selden (1956) and Latané (1960) rise and fall together with the opportunity cost variable; this, too, is as expected, since velocity is the reciprocal of the money-consumption ratio. Figure A1, in the appendix, confirms that very similar patterns appear when nominal GDP is used in place of nominal consumption spending in scaling the four Divisia aggregates and computing their velocities.

Figure 2, meanwhile, presents scatterplots that compare the behavior of the log Divisia money-consumption ratio or the consumption velocity of the Divisia monetary aggregate to either the log or the level of the user cost variable to look for preliminary evidence of the stable long-run relationship implied by each of three money demand specifications (24), (25), and (27). The color-coding in each panel provides signs of a possible shift in the parameters of each specification around a 1983:1 breakpoint that coincides with the end of the Volcker disinflation as well as the enactment of the 1982 Garn-St. Germain Act, which allowed banks to compete more effectively with money market mutual funds by issuing money market deposit accounts. The 1983 breakpoint also coincides with the termination of the sharp downward trends found in the money-consumption ratios and the correspondingly steep upward trends in consumption velocities shown in figure 1. From all panels in figure 2, it appears in particular that the responsiveness of Divisia money demand to changes in the user cost has diminished since 1983. Based on the indications from these graphs, estimates of the money demand relationships (24), (25) and (27) are reported below for both the full sample period starting in 1967:1 and the subsample beginning in 1983:1.

Examining each panel of figure 2 for evidence suggestive of a stable money demand function, the tightest linear relationships appear for Divisia M2 and MZM when the data transformations dictated by the Meltzer (1963) or Cagan (1956) specifications are applied. The plots for the Divisia M1 aggregate display signs of non-linearity at the lowest observed values of the opportunity cost that all three models (24), (25), and (27) would have difficulty accounting for. Finally, the relationship between Divisia money and its user costs looks weakest for the broad M4 aggregate, particularly in the post-1983 period. This casual inspection of graphs, however, cannot provide the same degree of precision as formal statistical results; hence, these are assembled next.

Unit Roots and Stationarity

Following Johansen and Juselius (1990), Ireland (2009), Judson, Schlusche, and Wong (2014), Benati, Lucas, Nicolini, and Weber (2016), Anderson, Bordo, and Duca (2017), Benati (2017) and many others, the theoretical money demand specifications (24), (25), and (27) are interpreted here as describing potential long-run cointegrating relationships between two variables: either the log money-consumption ratio or consumption velocity as a measure of the

real demand for money relative to a scale variable on the one hand, and either the log or the level of the user cost of money on the other. Taking this perspective requires that the two variables to be generated by time series processes contain unit roots.

Tables A1-A12 in the appendix, therefore, catalog results from Dickey-Fuller (1979) tests for unit roots in each of the variables appearing in (24), (25), and (27), augmented as in Said and Dickey (1979) to include q lagged first differences of the series. The constant term that would ordinarily appear in the Dickey-Fuller regression is removed, however, using the “local to unity GLS” detrending procedure that Elliot, Rothenberg, and Stock (1996) introduce to obtain an “efficient” form of the augmented Dickey-Fuller test that has greater power against stationary alternatives with largest autoregressive roots less than, but still very close to, one. Bootstrapped p -values for the ADF-GLS unit root test statistics, also reported in the tables, are computed as suggested by Benati (2015), by estimating ARIMA($q,1,0$) processes with the actual data on each variable, generating artificial data from the estimated ARIMA model 10,000 times, re-running the ADF-GLS statistic with each of the artificial data sets, and comparing the actual test statistic to the collection of statistics computed from the bootstrapped replications. By looking across all of the tables, one can see that the test results are robust to changing the level of monetary aggregation, the lag length q , the sample period, and the choice of consumer spending or aggregate income as scale variables for the monetary aggregates.

In particular, for all measures of $\ln(m_t^a)$ or v_t^a , the null hypothesis of a unit root in the level of the series is rejected, while the hypothesis of a unit root in the first difference is not. Clearly, the money-consumption, money-income, and consumption and income velocities of all four Divisia aggregates can be characterized as being driven by I(1) processes.

Curiously, however, most of the user cost measures are *not* characterized as nonstationary by the ADF-GLS tests. Whether in log form as $\ln(r_t - r_t^a)$ or in levels as $r_t - r_t^a$, the tests tend to reject their null hypothesis of a unit root in favor of the stationary alternative. The only exceptions appear in tables A2, A8, and A9 when the Divisia M1 user cost is considered. Looking back at figure 1, it becomes easy to understand the source of this result: not only do these user cost variables exhibit substantial transitory variation, but each passes through its mean repeatedly over the full sample.

On the other hand, visual inspection of the plots in figure 1 suggests that there are low-frequency trends in the user cost variables that mirror those in the scaled measures of Divisia money; these have, perhaps, been swamped by the more volatile transitory fluctuations in generating the unit root test results. If so, these nonstationary movements in the user costs will be picked up by tests for cointegration with the nonstationary Divisia monetary aggregates. This is exactly what happens in the results described next.

Tests for Cointegration

Tables 1-3 report full-sample estimates of the parameters of long-run money demand relationships of the forms proposed by Meltzer (1963), Cagan (1956), and Selden (1956) and Latané (1960). These estimates are obtained by normalizing the first (and, typically, the only) cointegrating vector found by Johansen's (1991) maximum likelihood approach, so that the long-run relationship takes the form shown in (24), (25), or (27), with a unitary coefficient on the log money-consumption ratio or the consumption velocity of Divisia money. In its unconstrained form, the vector autoregression within which these relationships appear includes q lags of each variable. But when constrained to feature a single cointegrating vector, this VAR(q) becomes an error-correction model that allows a constant term to enter into the cointegrating relationship but not into the remaining nonstationary component. As explained by Hamilton (1994, p.643), this additional restriction on the model's constant term rules out deterministic trends in the levels of the variables, to be consistent with the absence of any strong, uninterrupted upward or downward secular trend in the series plotted in figure 1.

Tables 1-3 also report the maximum eigenvalue test statistics of Johansen and Juselius (1990) and Johansen (1991), used first to test the null hypothesis of no ($r = 0$) cointegrating vector linking the scaled Divisia monetary aggregate and its user cost versus the alternative of one ($r = 1$) cointegrating vector and then the null hypothesis of one cointegrating vector versus the alternative of two. In both cases, the p -values shown are obtained via the bootstrapping procedures outlined by Cavaliere, Rahbek, and Taylor (2012) and recommended by Benati (2015) as well. For this application, 10,000 replications for the test statistic are generated from an estimated model that is "true" under the null hypothesis; the actual test statistic is then compared to the collection of bootstrapped replications. Finally, for each money demand specification and each level of monetary aggregation, results are obtained for values of q ranging from 2 through 5, with indications made for the optimal lag lengths – usually $q = 2$ but occasionally $q = 3$ – chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Starting with table 1, the Meltzer specification (24) characterizes the data well, particularly for the Divisia M1, M2, and MZM aggregates. The maximum eigenvalue tests reject their null hypothesis of no cointegration between the log money-consumption ratio and the log of the associated user cost, and the estimates of the parameter α_1 , measuring the absolute value of the constant elasticity of Divisia money demand with respect to the user cost, always have their expected, positive signs. Note that, given the strong evidence of nonstationarity in the log money-consumption ratio provided by the ADF-GLS tests in table A1, these tests providing equally strong evidence of cointegration also are consistent with the presence of stochastic trends in the Divisia user cost variables that, in table A2, the same unit root tests

miss. Benati (2017) relates similar findings linking low-frequency movements in simple-sum M1 and the three-month Treasury bill rate to Cochrane's (1994) time-series interpretation of the permanent income hypothesis. Just as in Cochrane's analysis, where near random-walk behavior in consumption works to disentangle permanent movements in real GDP from more volatile transitory fluctuations, the clear stochastic trend in the ratio of Divisia money to consumption helps find the unit root in the Divisia user cost that is hidden in a univariate setting by a much larger stationary component. In table 1, in fact, only the results for M4, which suggest the possibility of two cointegrating vectors between this broadest Divisia aggregate and its user cost, appear as a puzzle.

Table 2 repeats the analysis, with similar results, for the Cagan specification (25). Again, with the sole exception of Divisia M4, for which the tests find evidence of two cointegrating vectors, the results appear consistent with the theory. In particular, a single cointegrating vector linking the log money-consumption ratio to the level of the user cost appears for Divisia M1, M2, and MZM, with estimates of the constant semi-elasticity δ_1 always taking the expected sign.

Table 3 provides evidence of less satisfactory performance for the Selden-Latané specification (27). For Divisia M1, cointegration always appears between consumption velocity and its user cost, but estimates of the slope coefficient γ_1 appear highly sensitive to changes in the number of lags q included in the VAR; for the case where $q = 5$, this parameter even takes on the unexpected, negative sign. Results for this particular model become still worse when data on Divisia M4 are used: in this case, γ_1 is invariably negative. For Divisia M2 and MZM, however, the maximum eigenvalue tests continue to point to the existence of a single cointegrating vector, with estimates of γ_1 that are positive and stable as q increases from 2 to 5.

Tables A13-A15 repeat this full-sample analysis, with very similar results, using GDP in place of personal consumption expenditures as the scale variable in the money demand relationships. Thus the quarterly data from 1967:1 through 2017:2 appear consistent with the implications of the theory. Especially for M2 and MZM, stable long-run relationships clearly exist between the Divisia monetary aggregate and its user cost dual. And while Benati, Lucas, Nicolini, and Weber (2016) and Benati (2017) find that the Selden-Latané specification does better than other functional forms in describing the demand for simple-sum M1, the Meltzer

and Cagan models perform better here for the Divisia aggregates, again especially for M2 and MZM.⁶

In tables 4-6, where each of the models is re-estimated with data from the subsample beginning in 1983:1 and running through 2017:2, a number of additional results appear. First, the fit of all three money demand specifications deteriorates markedly for Divisia M1 and M4. For M1, estimates of the elasticity parameter α_1 from the Meltzer specification and the semi-elasticity parameter δ_1 from Cagan's model often have the unexpected, negative sign; estimates of γ_1 from the Selden-Latané specification are always positive, but vary greatly in magnitude as the lag length q of the VAR increases. For M4, estimates of these response parameters are always negative; moreover, the maximum eigenvalue tests often fail to reject their null of no cointegration between the Divisia aggregate and its user cost. Second, Divisia MZM also exhibits a slight deterioration in model performance; in particular, as the VAR lag length grows, the evidence favoring the existence of a cointegrating money demand relation weakens.

But third and most important, for Divisia M2 the results remain strong across the board. The maximum eigenvalue tests point consistently to the existence of a single cointegrating relation between Divisia M2 and its user cost, with elasticity parameters that are stable and have the expected sign. Fourth and finally, estimates of all three models reveal that the responsiveness of Divisia M2 demand to changes in its user cost dual have declined in magnitude since 1983, compared to the full sample. This last result confirms the visual impression provided earlier by figure 2, where the slope of the scatterplots comparing Divisia M2 to its user cost flatten out after the 1983 breakpoint.

Carl Christ (1993) argues that the most persuasive evidence supporting any econometric model comes from its ability to fit data that were generated after the specification was initially formulated.⁷ He suggests (p.73), "that what we economists should do is formulate our models, then go fishing for 50 years and let new data accumulate, and finally come back and confront our models with the new data." More than 50 years ago, in fact, Meltzer (1963) estimated a money demand specification of the form shown here in (24). With annual data, 1900-1958, he obtained (p.225, equation 3') an estimate of -0.50 for the interest elasticity of M2 demand. In addition to using the stock of wealth as opposed to a flow of spending or

⁶ It also might be noted that the finding by Friedman and Kuttner (1992) of no stable money demand relationship for simple sum M1 or M2 in their influential paper was reversed when it was replicated by Hendrickson (2014) with Divisia measures of the same aggregates.

⁷ Lucas (1988) makes a similar point, with direct reference to Meltzer's (1963) model of money demand.

income as his scale variable, Meltzer used a simple-sum M2 aggregate together with a long-term corporate bond rate to measure its opportunity cost. If one interprets these measures as error-ridden, relative to the Divisia M2 aggregate and its associated user cost, one would expect the resulting elasticity estimates to be biased towards zero. And, indeed, the elasticity estimates obtained here, from the same model but different data, are larger in magnitude: between -1.6 and -1.8 as shown in table 1 for the full 1967-2017 sample and between -0.70 and -0.85 in table 4 for the post-1983 subsample.

Alternatively, one can accept that institutional changes within the United States financial system have led to changes in the elasticity of money demand over time, even as the general money demand specification proposed originally by Meltzer (1963) continues to fit the data well, without the need for any additional explanatory or dummy variables. In that case, however, it is noteworthy that the largest elasticity estimates across those reported here and in Meltzer's original study appear when the model is estimated using data from the high inflation years of the 1970s and early 1980s. This finding provides another reason to be skeptical of popular discussions, associating historical periods of low inflation – the Great Depression of the late 1920s and 1930s and, more recently, the period of zero nominal interest rates following the financial crisis and Great Recession of 2007-09 – with a Keynesian liquidity trap since, for a liquidity trap to take hold, the demand for money ought to become *more* elastic as inflation falls.⁸

Conclusions and Implications

The economic theory presented here suggests that, if stable long-run money demand relationships exist, they will most likely serve to link the demand for a Divisia monetary aggregate to its user cost dual. The accompanying statistical results show, in fact, that relations of precisely this form appear in quarterly United States data, particularly for the Divisia M2 aggregate.

Importantly, the data used to estimate these stable demand relationships for Divisia M2 include observations from periods before, during, and since the financial crisis and Great Recession of 2007-09 and the extended period of zero short-term interest rates that followed. Thus, these results complement those presented earlier by Keating, Kelly, Smith, and Valcarcel (2014) and Belongia and Ireland (2015b, 2016, 2017, 2018), which show that information in the Divisia monetary aggregates can be useful in gauging the stance of monetary policy and estimating the effects that monetary policy has on output and inflation. Indeed, when traditional interest rate policies are constrained during exceptional periods by the by the zero

⁸ Along similar lines, it should also be noted that Keynes (1936, p.207) himself expressed reservations about the empirical relevance of the liquidity trap, explaining that, “whilst this limiting case might become practically important in future, I know of no example of it hitherto.”

lower bound, the existence of a stable demand for money function offers an alternative approach to monetary policy based on targeting a path for the quantity of money. Moreover, even during more normal periods of interest rates, the existence of a stable demand function for Divisia M2 indicates the price level or inflation rate could be stabilized by controlling the behavior of this aggregate. After years of neglect, the existence of a stable demand for money function motivates a reconsideration of quantity-theoretic approaches to monetary policymaking.

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Table 1. Cointegrating Money Demand RelationshipsMeltzer Specification: $\ln(m_t^a) = \alpha_0 - \alpha_1 \ln(r_t - r_t^a)$ $\ln(m_t^a)$: log money-consumption ratio

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | α_0 | α_1 | Max Eigenvalue $r = 0$ p -value | | Max Eigenvalue $r = 1$ p -value | |
|----------------------|----------------|------------|------------|--------------------------------------|--------|--------------------------------------|--------|
| M1 | 2 ⁺ | -8.53 | 2.75 | 46.36 | 0.0001 | 3.79 | 0.5466 |
| | 3 | -7.35 | 2.24 | 37.35 | 0.0001 | 4.11 | 0.4988 |
| | 4 | -6.44 | 1.85 | 27.69 | 0.0009 | 3.28 | 0.6337 |
| | 5 | -6.68 | 1.96 | 20.79 | 0.0156 | 2.62 | 0.7463 |
| M2 | 2 ⁺ | -4.93 | 1.79 | 48.28 | 0.0001 | 6.32 | 0.1979 |
| | 3 | -4.84 | 1.72 | 44.91 | 0.0001 | 6.31 | 0.2145 |
| | 4 | -4.71 | 1.63 | 36.39 | 0.0001 | 6.63 | 0.1858 |
| | 5 | -4.92 | 1.77 | 33.83 | 0.0002 | 5.11 | 0.3157 |
| MZM | 2 ⁺ | -4.59 | 1.57 | 42.23 | 0.0001 | 7.01 | 0.1385 |
| | 3 | -4.43 | 1.45 | 40.37 | 0.0001 | 6.87 | 0.1558 |
| | 4 | -4.43 | 1.45 | 29.54 | 0.0001 | 7.25 | 0.1375 |
| | 5 | -4.47 | 1.47 | 29.65 | 0.0005 | 5.71 | 0.2499 |
| M4 | 2 ⁺ | -9.99 | 5.82 | 28.29 | 0.0005 | 9.71 | 0.0427 |
| | 3 | -8.14 | 4.39 | 24.81 | 0.0028 | 10.32 | 0.0364 |
| | 4 | -11.05 | 6.55 | 19.28 | 0.0192 | 10.41 | 0.0293 |
| | 5 | -8.65 | 4.77 | 17.41 | 0.0403 | 8.62 | 0.0720 |

Notes: The log money-consumption ratio is computed by dividing the indicated monetary aggregate by nominal personal consumption expenditures and taking the natural logarithm. α_0 and α_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table 2. Cointegrating Money Demand RelationshipsCagan Specification: $\ln(m_t^a) = \delta_0 - \delta_1(r_t - r_t^a)$ $\ln(m_t^a)$: log money-consumption ratio

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | δ_0 | δ_1 | Max Eigenvalue | | Max Eigenvalue | |
|----------------------|-----------------|------------|------------|----------------|------------|----------------|------------|
| | | | | $r = 0$ | p -value | $r = 1$ | p -value |
| M1 | 2 ⁺⁺ | 1.81 | 33.56 | 32.18 | 0.0004 | 3.13 | 0.6032 |
| | 3 | 1.20 | 28.57 | 25.90 | 0.0019 | 4.10 | 0.4755 |
| | 4 | 0.50 | 22.58 | 17.74 | 0.0419 | 3.27 | 0.6046 |
| | 5 | 3.17 | 45.54 | 16.61 | 0.0620 | 2.55 | 0.7216 |
| M2 | 2 ⁺⁺ | -0.75 | 6.17 | 45.55 | 0.0001 | 6.39 | 0.1886 |
| | 3 | -0.74 | 6.21 | 41.89 | 0.0001 | 6.33 | 0.2066 |
| | 4 | -0.85 | 5.84 | 33.74 | 0.0002 | 6.78 | 0.1689 |
| | 5 | -0.55 | 6.95 | 35.47 | 0.0001 | 5.49 | 0.2613 |
| MZM | 2 ⁺⁺ | -0.99 | 5.26 | 37.64 | 0.0001 | 7.05 | 0.1372 |
| | 3 ⁺ | -0.98 | 5.29 | 39.02 | 0.0001 | 6.81 | 0.1635 |
| | 4 | -1.05 | 5.06 | 27.15 | 0.0014 | 7.16 | 0.1363 |
| | 5 | -0.89 | 5.66 | 31.39 | 0.0003 | 5.83 | 0.2323 |
| M4 | 2 ⁺⁺ | 1.33 | 13.16 | 26.00 | 0.0021 | 9.31 | 0.0534 |
| | 3 | 0.40 | 9.84 | 23.64 | 0.0045 | 9.55 | 0.0506 |
| | 4 | 0.77 | 11.33 | 17.37 | 0.0391 | 9.93 | 0.0366 |
| | 5 | 0.97 | 12.02 | 18.34 | 0.0298 | 8.22 | 0.0800 |

Notes: The log money-consumption ratio is computed by dividing the indicated monetary aggregate by nominal personal consumption expenditures and taking the natural logarithm. δ_0 and δ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table 3. Cointegrating Money Demand RelationshipsSelden-Latané Specification: $v_t^a = \gamma_0 + \gamma_1(r_t - r_t^a)$ v_t^a : consumption velocity

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | γ_0 | γ_1 | Max Eigenvalue | | Max Eigenvalue | |
|----------------------|-----------------|------------|------------|----------------|------------|----------------|------------|
| | | | | $r = 0$ | p -value | $r = 1$ | p -value |
| M1 | 2 ⁺⁺ | -175.73 | 1581.36 | 27.32 | 0.0014 | 3.17 | 0.6006 |
| | 3 | -655.66 | 5775.83 | 22.10 | 0.0079 | 4.27 | 0.4463 |
| | 4 | -58.37 | 580.16 | 16.49 | 0.0521 | 3.20 | 0.6001 |
| | 5 | 2278.97 | -19762.76 | 16.20 | 0.0677 | 2.68 | 0.6973 |
| M2 | 2 ⁺⁺ | -13.71 | 92.36 | 35.52 | 0.0001 | 4.50 | 0.3642 |
| | 3 | -13.98 | 93.68 | 33.58 | 0.0001 | 4.53 | 0.3869 |
| | 4 | -13.27 | 91.59 | 27.93 | 0.0013 | 4.66 | 0.3506 |
| | 5 | -16.43 | 103.09 | 30.24 | 0.0006 | 4.02 | 0.4429 |
| MZM | 2 [*] | -10.98 | 82.31 | 29.48 | 0.0008 | 4.73 | 0.3305 |
| | 3 ⁺ | -11.47 | 84.78 | 30.33 | 0.0003 | 4.96 | 0.3134 |
| | 4 | -12.18 | 88.03 | 21.59 | 0.0092 | 4.95 | 0.3111 |
| | 5 | -12.03 | 86.81 | 25.35 | 0.0022 | 4.25 | 0.3945 |
| M4 | 2 ⁺⁺ | 228.01 | -814.56 | 23.13 | 0.0050 | 7.41 | 0.1122 |
| | 3 | 539.02 | -2002.32 | 20.62 | 0.0108 | 7.51 | 0.1078 |
| | 4 | 57.63 | -183.79 | 15.66 | 0.0656 | 7.59 | 0.1087 |
| | 5 | 129.64 | -453.79 | 16.33 | 0.0575 | 6.97 | 0.1379 |

Notes: Consumption velocity is computed by dividing nominal personal consumption expenditures by the indicated monetary aggregate. γ_0 and γ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on v_t^a equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table 4. Cointegrating Money Demand RelationshipsMeltzer Specification: $\ln(m_t^a) = \alpha_0 - \alpha_1 \ln(r_t - r_t^a)$ $\ln(m_t^a)$: log money-consumption ratio

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | α_0 | α_1 | Max Eigenvalue | | Max Eigenvalue | |
|----------------------|----------------|------------|------------|----------------|------------|----------------|------------|
| | | | | $r = 0$ | p -value | $r = 1$ | p -value |
| M1 | 2 ⁺ | -7.30 | 2.18 | 31.13 | 0.0008 | 2.47 | 0.7902 |
| | 3 | -10.31 | 3.46 | 24.89 | 0.0029 | 2.02 | 0.8526 |
| | 4 | 88.11 | 38.55 | 20.43 | 0.0173 | 1.50 | 0.9355 |
| | 5 | 58.99 | -26.18 | 14.59 | 0.1203 | 1.70 | 0.9136 |
| M2 | 2 ⁺ | -3.37 | 0.70 | 26.75 | 0.0008 | 0.81 | 0.9828 |
| | 3 | -3.42 | 0.73 | 25.53 | 0.0028 | 0.72 | 0.9887 |
| | 4 | -3.50 | 0.79 | 22.44 | 0.0074 | 0.44 | 0.9994 |
| | 5 | -3.57 | 0.84 | 18.89 | 0.0263 | 0.39 | 0.9997 |
| MZM | 2 ⁺ | -3.34 | 0.66 | 27.49 | 0.0010 | 0.87 | 0.9718 |
| | 3 | -3.37 | 0.68 | 20.95 | 0.0114 | 0.96 | 0.9634 |
| | 4 | -3.40 | 0.71 | 17.41 | 0.0421 | 0.91 | 0.9712 |
| | 5 | -3.49 | 0.77 | 13.32 | 0.1523 | 1.49 | 0.8961 |
| M4 | 2 ⁺ | -1.18 | -0.76 | 15.94 | 0.0598 | 6.55 | 0.1517 |
| | 3 | -1.33 | -0.65 | 14.50 | 0.1047 | 5.67 | 0.2201 |
| | 4 | -1.61 | -0.44 | 13.17 | 0.1683 | 5.23 | 0.2649 |
| | 5 | -1.79 | -0.31 | 11.45 | 0.2698 | 4.90 | 0.2892 |

Notes: The log money-consumption ratio is computed by dividing the indicated monetary aggregate by nominal personal consumption expenditures and taking the natural logarithm. α_0 and α_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table 5. Cointegrating Money Demand RelationshipsCagan Specification: $\ln(m_t^a) = \delta_0 - \delta_1(r_t - r_t^a)$ $\ln(m_t^a)$: log money-consumption ratio

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | δ_0 | δ_1 | Max Eigenvalue | | Max Eigenvalue | |
|----------------------|-----------------|------------|------------|----------------|------------|----------------|------------|
| | | | | $r = 0$ | p -value | $r = 1$ | p -value |
| M1 | 2 ⁺⁺ | 9.83 | 111.84 | 29.48 | 0.0007 | 1.59 | 0.9109 |
| | 3 | -45.34 | -403.72 | 25.21 | 0.0034 | 1.87 | 0.8725 |
| | 4 | -4.92 | -25.49 | 21.08 | 0.0133 | 1.56 | 0.9163 |
| | 5 | -5.01 | -26.37 | 17.51 | 0.0490 | 1.88 | 0.8750 |
| M2 | 2 ⁺⁺ | -1.54 | 3.23 | 24.51 | 0.0044 | 0.70 | 0.9880 |
| | 3 | -1.54 | 3.24 | 22.45 | 0.0079 | 0.58 | 0.9943 |
| | 4 | -1.45 | 3.56 | 19.88 | 0.0182 | 0.34 | 0.9997 |
| | 5 | -1.43 | 3.63 | 19.64 | 0.0224 | 0.32 | 0.9998 |
| MZM | 2 ⁺⁺ | -1.57 | 3.21 | 25.19 | 0.0033 | 0.86 | 0.9740 |
| | 3 | -1.57 | 3.22 | 20.14 | 0.0140 | 0.94 | 0.9710 |
| | 4 | -1.52 | 3.41 | 17.27 | 0.0441 | 0.95 | 0.9642 |
| | 5 | -1.50 | 3.50 | 15.35 | 0.0824 | 1.82 | 0.8246 |
| M4 | 2 ⁺⁺ | -2.88 | -2.50 | 15.40 | 0.0691 | 6.51 | 0.1545 |
| | 3 | -2.79 | -2.19 | 14.65 | 0.0912 | 5.60 | 0.2207 |
| | 4 | -2.59 | -1.44 | 13.06 | 0.1654 | 5.16 | 0.2695 |
| | 5 | -2.53 | -1.18 | 13.14 | 0.1696 | 4.92 | 0.2978 |

Notes: The log money-consumption ratio is computed by dividing the indicated monetary aggregate by nominal personal consumption expenditures and taking the natural logarithm. δ_0 and δ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table 6. Cointegrating Money Demand RelationshipsSelden-Latané Specification: $v_t^a = \gamma_0 + \gamma_1(r_t - r_t^a)$ v_t^a : consumption velocity

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | γ_0 | γ_1 | Max Eigenvalue | | Max Eigenvalue | |
|----------------------|-----------------|------------|------------|----------------|------------|----------------|------------|
| | | | | $r = 0$ | p -value | $r = 1$ | p -value |
| M1 | 2 ⁺⁺ | -8.03 | 106.17 | 26.77 | 0.0018 | 2.06 | 0.8316 |
| | 3 | -10.69 | 186.99 | 21.91 | 0.0085 | 1.89 | 0.8652 |
| | 4 | -60.27 | 655.67 | 18.66 | 0.0285 | 1.31 | 0.9484 |
| | 5 | -64.94 | 700.98 | 16.97 | 0.0519 | 1.48 | 0.9264 |
| M2 | 2 ⁺⁺ | 2.62 | 31.78 | 24.17 | 0.0032 | 0.88 | 0.9768 |
| | 3 | 2.68 | 31.64 | 22.19 | 0.0083 | 0.67 | 0.9903 |
| | 4 | 2.00 | 34.23 | 19.64 | 0.0190 | 0.36 | 0.9993 |
| | 5 | 1.89 | 34.62 | 19.48 | 0.0246 | 0.28 | 0.9999 |
| MZM | 2 ⁺⁺ | 2.98 | 31.48 | 25.09 | 0.0028 | 1.10 | 0.9512 |
| | 3 | 3.03 | 31.45 | 20.07 | 0.0170 | 0.97 | 0.9634 |
| | 4 | 2.53 | 33.46 | 16.94 | 0.0549 | 0.72 | 0.9871 |
| | 5 | 2.52 | 33.39 | 14.69 | 0.1005 | 1.05 | 0.9476 |
| M4 | 2 ⁺⁺ | 15.48 | -23.65 | 15.29 | 0.0750 | 6.19 | 0.1784 |
| | 3 | 14.36 | -19.69 | 14.77 | 0.0956 | 5.42 | 0.2410 |
| | 4 | 12.60 | -13.03 | 13.06 | 0.1663 | 5.05 | 0.2757 |
| | 5 | 12.07 | -10.80 | 12.88 | 0.1763 | 4.83 | 0.3071 |

Notes: Consumption velocity is computed by dividing nominal personal consumption expenditures by the indicated monetary aggregate. γ_0 and γ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on v_t^a equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.



Figure 1. Divisia Monetary Data. The log money-consumption ratio in column one and consumption velocity in column four are computed using nominal personal consumption expenditures as the scale variable.

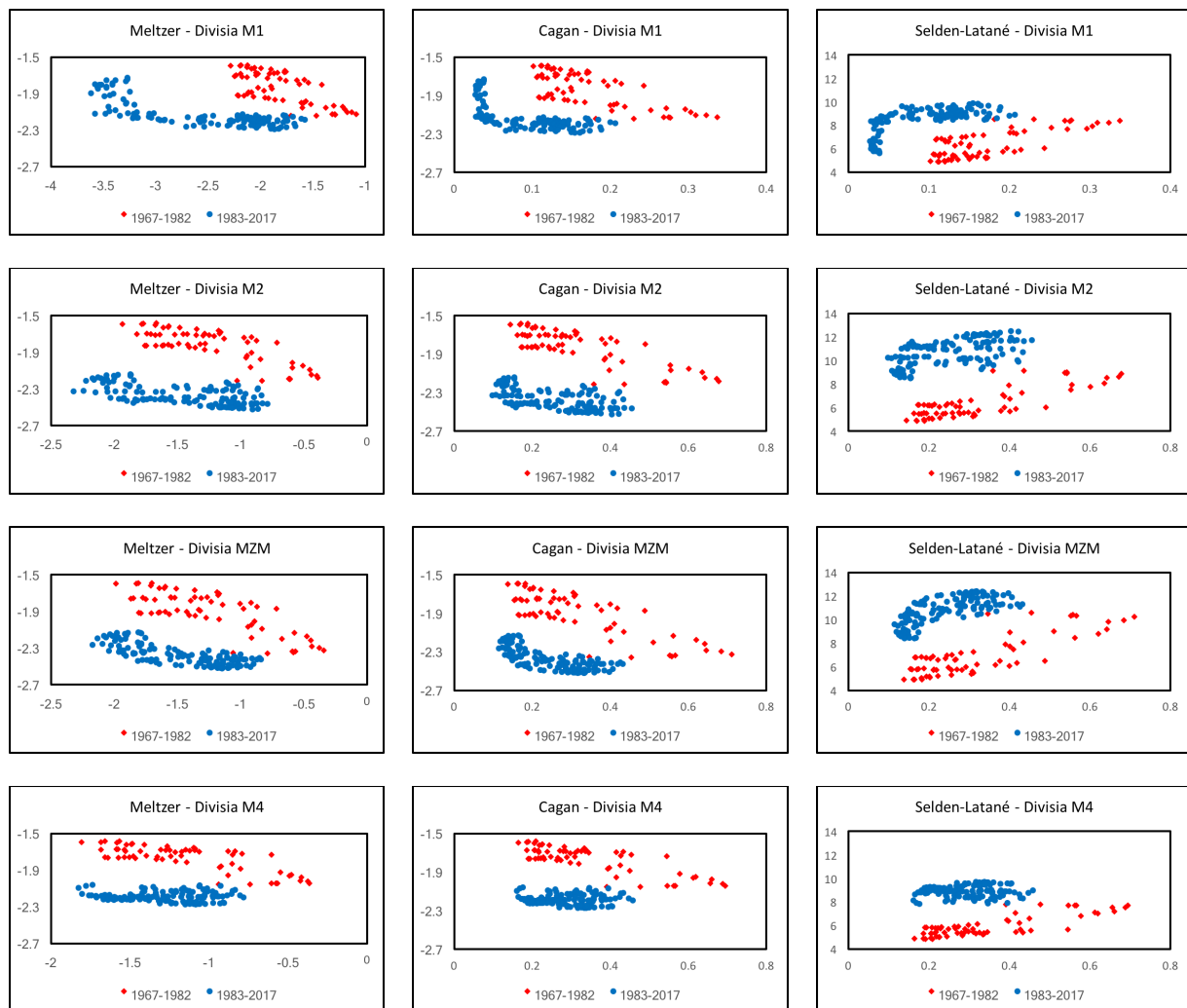


Figure 2. Divisia Money Demand Relationships. Each scatterplot compares the indicated Divisia monetary aggregate, scaled by nominal personal consumption expenditures, measured along the vertical axis, to the associated user cost aggregate, measured along the horizontal axis. The Meltzer (1963) specification relates the log money-consumption ratio to the log of the user cost; the Cagan (1956) specification relates the log money-consumption ratio to the level of the user cost; the Selden-Latané specification relates the consumption velocity of money to the level of the user cost.

Table A1. Efficient Unit Root Test Results $\ln(m_t^a)$: log money-consumption ratio

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.49 | 0.5627 | -5.13 | 0.0001 |
| | 2 | -0.58 | 0.5289 | -4.16 | 0.0001 |
| | 3 | -0.64 | 0.5044 | -3.66 | 0.0001 |
| | 4 | -0.66 | 0.4923 | -3.34 | 0.0010 |
| M2 | 1 | -0.01 | 0.3783 | -5.40 | 0.0001 |
| | 2 | -0.11 | 0.4028 | -4.46 | 0.0002 |
| | 3 | -0.20 | 0.4262 | -4.13 | 0.0002 |
| | 4 | -0.18 | 0.4334 | -3.61 | 0.0002 |
| MZM | 1 | -0.19 | 0.4446 | -5.56 | 0.0001 |
| | 2 | -0.22 | 0.4425 | -4.41 | 0.0001 |
| | 3 | -0.32 | 0.4718 | -4.44 | 0.0001 |
| | 4 | -0.24 | 0.4601 | -3.85 | 0.0002 |
| M4 | 1 | 0.64 | 0.3061 | -6.47 | 0.0001 |
| | 2 | 0.52 | 0.3051 | -5.36 | 0.0001 |
| | 3 | 0.39 | 0.3250 | -5.24 | 0.0001 |
| | 4 | 0.49 | 0.3194 | -4.64 | 0.0001 |

Notes: The log money-consumption ratio is computed by dividing the indicated monetary aggregate by nominal personal consumption expenditures and taking the natural logarithm. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A2. Efficient Unit Root Test Results

$\ln(r_t - r_t^a)$: log user cost
Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -1.41 | 0.1620 | -7.84 | 0.0001 |
| | 2 | -1.31 | 0.1836 | -5.65 | 0.0001 |
| | 3 | -1.49 | 0.1372 | -4.55 | 0.0001 |
| | 4 | -1.60 | 0.1104 | -3.59 | 0.0005 |
| M2 | 1 | -2.24 | 0.0320 | -6.39 | 0.0001 |
| | 2 | -2.12 | 0.0421 | -4.66 | 0.0001 |
| | 3 | -2.15 | 0.0385 | -3.64 | 0.0003 |
| | 4 | -2.21 | 0.0314 | -2.82 | 0.0066 |
| MZM | 1 | -2.28 | 0.0257 | -6.38 | 0.0001 |
| | 2 | -2.08 | 0.0436 | -4.52 | 0.0001 |
| | 3 | -2.12 | 0.0363 | -3.44 | 0.0009 |
| | 4 | -2.23 | 0.0287 | -2.68 | 0.0067 |
| M4 | 1 | -2.33 | 0.0226 | -7.65 | 0.0001 |
| | 2 | -2.10 | 0.0434 | -5.57 | 0.0001 |
| | 3 | -1.99 | 0.0542 | -4.20 | 0.0001 |
| | 4 | -2.02 | 0.0528 | -3.20 | 0.0015 |

Notes: ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A3. Efficient Unit Root Test Results $r_t - r_t^a$: user cost

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -2.17 | 0.0346 | -9.32 | 0.0001 |
| | 2 | -1.70 | 0.0968 | -6.15 | 0.0001 |
| | 3 | -1.97 | 0.0536 | -4.85 | 0.0001 |
| | 4 | -2.14 | 0.0388 | -3.68 | 0.0002 |
| M2 | 1 | -2.41 | 0.0191 | -8.72 | 0.0001 |
| | 2 | -2.18 | 0.0359 | -6.44 | 0.0001 |
| | 3 | -2.25 | 0.0278 | -4.88 | 0.0001 |
| | 4 | -2.47 | 0.0145 | -3.82 | 0.0001 |
| MZM | 1 | -2.50 | 0.0164 | -9.31 | 0.0001 |
| | 2 | -2.11 | 0.0412 | -6.48 | 0.0001 |
| | 3 | -2.23 | 0.0287 | -4.95 | 0.0001 |
| | 4 | -2.41 | 0.0195 | -3.81 | 0.0001 |
| M4 | 1 | -2.49 | 0.0159 | -9.87 | 0.0001 |
| | 2 | -2.24 | 0.0273 | -7.67 | 0.0001 |
| | 3 | -2.09 | 0.0414 | -5.67 | 0.0001 |
| | 4 | -2.23 | 0.0295 | -4.26 | 0.0001 |

Notes: ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A4. Efficient Unit Root Test Results v_t^a : consumption velocity

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.59 | 0.5248 | -5.27 | 0.0001 |
| | 2 | -0.67 | 0.4908 | -4.60 | 0.0001 |
| | 3 | -0.70 | 0.4749 | -4.20 | 0.0001 |
| | 4 | -0.70 | 0.4714 | -3.92 | 0.0001 |
| M2 | 1 | -0.23 | 0.4682 | -5.59 | 0.0001 |
| | 2 | -0.29 | 0.4712 | -4.84 | 0.0001 |
| | 3 | -0.34 | 0.4667 | -4.54 | 0.0001 |
| | 4 | -0.32 | 0.4719 | -3.98 | 0.0001 |
| MZM | 1 | -0.43 | 0.4835 | -5.66 | 0.0001 |
| | 2 | -0.43 | 0.4887 | -4.68 | 0.0001 |
| | 3 | -0.51 | 0.4862 | -4.81 | 0.0001 |
| | 4 | -0.43 | 0.4832 | -4.23 | 0.0001 |
| M4 | 1 | 0.50 | 0.3501 | -6.45 | 0.0001 |
| | 2 | 0.39 | 0.3534 | -5.48 | 0.0001 |
| | 3 | 0.30 | 0.3707 | -5.43 | 0.0001 |
| | 4 | 0.40 | 0.3654 | -4.88 | 0.0001 |

Notes: Consumption velocity is computed by dividing nominal personal consumption expenditures by the indicated monetary aggregate. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A5. Efficient Unit Root Test Results $\ln(m_t^a)$: log money-output ratio

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.51 | 0.5685 | -5.12 | 0.0001 |
| | 2 | -0.59 | 0.5342 | -3.97 | 0.0001 |
| | 3 | -0.69 | 0.4844 | -3.76 | 0.0001 |
| | 4 | -0.67 | 0.4900 | -3.38 | 0.0003 |
| M2 | 1 | -0.22 | 0.4469 | -4.37 | 0.0001 |
| | 2 | -0.30 | 0.4620 | -3.35 | 0.0014 |
| | 3 | -0.39 | 0.4865 | -3.22 | 0.0018 |
| | 4 | -0.35 | 0.4778 | -2.85 | 0.0053 |
| MZM | 1 | -0.37 | 0.4894 | -5.19 | 0.0001 |
| | 2 | -0.40 | 0.4804 | -4.03 | 0.0001 |
| | 3 | -0.49 | 0.4812 | -4.10 | 0.0001 |
| | 4 | -0.42 | 0.4910 | -3.57 | 0.0006 |
| M4 | 1 | 0.20 | 0.3793 | -4.64 | 0.0001 |
| | 2 | 0.12 | 0.3874 | -3.49 | 0.0009 |
| | 3 | -0.02 | 0.4197 | -3.48 | 0.0005 |
| | 4 | 0.11 | 0.4070 | -3.07 | 0.0028 |

Notes: The log money-output ratio is computed by dividing the indicated monetary aggregate by nominal GDP and taking the natural logarithm. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A6. Efficient Unit Root Test Results v_t^a : income velocity

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.61 | 0.5381 | -5.40 | 0.0001 |
| | 2 | -0.69 | 0.4827 | -4.36 | 0.0001 |
| | 3 | -0.78 | 0.4441 | -4.42 | 0.0001 |
| | 4 | -0.71 | 0.4817 | -3.99 | 0.0001 |
| M2 | 1 | -0.39 | 0.4995 | -5.12 | 0.0001 |
| | 2 | -0.45 | 0.5022 | -4.16 | 0.0001 |
| | 3 | -0.52 | 0.4947 | -4.16 | 0.0001 |
| | 4 | -0.46 | 0.5017 | -3.65 | 0.0001 |
| MZM | 1 | -0.55 | 0.4950 | -5.63 | 0.0001 |
| | 2 | -0.58 | 0.4885 | -4.50 | 0.0001 |
| | 3 | -0.66 | 0.4666 | -4.83 | 0.0001 |
| | 4 | -0.56 | 0.4912 | -4.11 | 0.0001 |
| M4 | 1 | 0.04 | 0.4164 | -5.34 | 0.0001 |
| | 2 | -0.03 | 0.4270 | -4.10 | 0.0002 |
| | 3 | -0.16 | 0.4363 | -4.26 | 0.0001 |
| | 4 | 0.00 | 0.4229 | -3.71 | 0.0002 |

Notes: Income velocity is computed by dividing nominal GDP by the indicated monetary aggregate. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A7. Efficient Unit Root Test Results $\ln(m_t^a)$: log money-consumption ratio

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | 0.55 | 0.6858 | -4.28 | 0.0001 |
| | 2 | 0.40 | 0.6470 | -3.63 | 0.0002 |
| | 3 | 0.18 | 0.6046 | -3.14 | 0.0014 |
| | 4 | 0.01 | 0.5680 | -3.09 | 0.0032 |
| M2 | 1 | -0.70 | 0.5001 | -3.06 | 0.0031 |
| | 2 | -0.71 | 0.4986 | -2.65 | 0.0115 |
| | 3 | -0.70 | 0.5150 | -2.32 | 0.0249 |
| | 4 | -0.70 | 0.5097 | -1.96 | 0.0589 |
| MZM | 1 | -1.15 | 0.2605 | -4.42 | 0.0001 |
| | 2 | -0.86 | 0.3665 | -3.86 | 0.0001 |
| | 3 | -0.93 | 0.3350 | -3.96 | 0.0001 |
| | 4 | -0.66 | 0.4353 | -3.39 | 0.0010 |
| M4 | 1 | -0.85 | 0.3574 | -4.10 | 0.0001 |
| | 2 | -0.83 | 0.3708 | -3.50 | 0.0007 |
| | 3 | -0.82 | 0.3713 | -3.39 | 0.0008 |
| | 4 | -0.66 | 0.4200 | -2.86 | 0.0052 |

Notes: The log money-consumption ratio is computed by dividing the indicated monetary aggregate by nominal personal consumption expenditures and taking the natural logarithm. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A8. Efficient Unit Root Test Results

$\ln(r_t - r_t^a)$: log user cost
Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.65 | 0.4098 | -7.49 | 0.0001 |
| | 2 | -0.73 | 0.3861 | -5.65 | 0.0001 |
| | 3 | -0.89 | 0.3301 | -4.99 | 0.0001 |
| | 4 | -0.90 | 0.3341 | -4.11 | 0.0003 |
| M2 | 1 | -1.72 | 0.1063 | -8.19 | 0.0001 |
| | 2 | -1.76 | 0.0899 | -6.39 | 0.0001 |
| | 3 | -1.85 | 0.0731 | -5.64 | 0.0001 |
| | 4 | -1.82 | 0.0776 | -4.66 | 0.0001 |
| MZM | 1 | -1.86 | 0.0751 | -8.36 | 0.0001 |
| | 2 | -1.86 | 0.0725 | -6.42 | 0.0001 |
| | 3 | -1.93 | 0.0611 | -5.54 | 0.0001 |
| | 4 | -1.92 | 0.0611 | -4.54 | 0.0001 |
| M4 | 1 | -2.71 | 0.0091 | -10.18 | 0.0001 |
| | 2 | -2.41 | 0.0184 | -8.02 | 0.0001 |
| | 3 | -2.30 | 0.0243 | -6.77 | 0.0001 |
| | 4 | -2.24 | 0.0314 | -5.59 | 0.0001 |

Notes: ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A9. Efficient Unit Root Test Results $r_t - r_t^a$: user cost

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.83 | 0.3650 | -6.59 | 0.0001 |
| | 2 | -0.89 | 0.3429 | -5.14 | 0.0001 |
| | 3 | -0.98 | 0.3105 | -4.31 | 0.0001 |
| | 4 | -1.10 | 0.2574 | -3.94 | 0.0001 |
| M2 | 1 | -1.94 | 0.0668 | -7.57 | 0.0001 |
| | 2 | -2.05 | 0.0473 | -5.99 | 0.0001 |
| | 3 | -2.18 | 0.0363 | -4.97 | 0.0001 |
| | 4 | -2.37 | 0.0224 | -4.42 | 0.0001 |
| MZM | 1 | -2.02 | 0.0540 | -7.83 | 0.0001 |
| | 2 | -2.05 | 0.0493 | -6.08 | 0.0001 |
| | 3 | -2.11 | 0.0410 | -4.90 | 0.0001 |
| | 4 | -2.29 | 0.0265 | -4.19 | 0.0001 |
| M4 | 1 | -2.73 | 0.0073 | -9.99 | 0.0001 |
| | 2 | -2.47 | 0.0188 | -8.13 | 0.0001 |
| | 3 | -2.30 | 0.0261 | -6.56 | 0.0001 |
| | 4 | -2.35 | 0.0215 | -5.36 | 0.0001 |

Notes: ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A10. Efficient Unit Root Test Results v_t^a : consumption velocity

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | -0.07 | 0.5419 | -4.25 | 0.0001 |
| | 2 | -0.23 | 0.5165 | -3.84 | 0.0001 |
| | 3 | -0.28 | 0.4970 | -3.46 | 0.0005 |
| | 4 | -0.34 | 0.4889 | -3.34 | 0.0011 |
| M2 | 1 | -0.78 | 0.4672 | -3.36 | 0.0015 |
| | 2 | -0.78 | 0.4744 | -2.93 | 0.0049 |
| | 3 | -0.77 | 0.4852 | -2.65 | 0.0091 |
| | 4 | -0.74 | 0.4942 | -2.24 | 0.0331 |
| MZM | 1 | -1.33 | 0.2081 | -4.36 | 0.0001 |
| | 2 | -1.10 | 0.2893 | -3.87 | 0.0003 |
| | 3 | -1.14 | 0.2713 | -4.03 | 0.0001 |
| | 4 | -0.86 | 0.3787 | -3.46 | 0.0005 |
| M4 | 1 | -0.86 | 0.3588 | -4.20 | 0.0001 |
| | 2 | -0.86 | 0.3679 | -3.59 | 0.0004 |
| | 3 | -0.86 | 0.3642 | -3.51 | 0.0006 |
| | 4 | -0.70 | 0.3997 | -3.02 | 0.0038 |

Notes: Consumption velocity is computed by dividing nominal personal consumption expenditures by the indicated monetary aggregate. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A11. Efficient Unit Root Test Results $\ln(m_t^a)$: log money-output ratio

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | 1.00 | 0.7657 | -4.27 | 0.0001 |
| | 2 | 0.91 | 0.7560 | -3.57 | 0.0007 |
| | 3 | 0.64 | 0.7042 | -3.18 | 0.0017 |
| | 4 | 0.52 | 0.6755 | -3.11 | 0.0026 |
| M2 | 1 | -0.62 | 0.4923 | -2.88 | 0.0058 |
| | 2 | -0.63 | 0.4855 | -2.67 | 0.0100 |
| | 3 | -0.54 | 0.5239 | -2.22 | 0.0322 |
| | 4 | -0.61 | 0.5067 | -2.00 | 0.0530 |
| MZM | 1 | -0.89 | 0.3485 | -4.29 | 0.0001 |
| | 2 | -0.52 | 0.4578 | -3.99 | 0.0002 |
| | 3 | -0.46 | 0.4663 | -3.64 | 0.0008 |
| | 4 | -0.44 | 0.4674 | -3.36 | 0.0011 |
| M4 | 1 | -1.63 | 0.1288 | -3.89 | 0.0005 |
| | 2 | -1.54 | 0.1514 | -3.32 | 0.0009 |
| | 3 | -1.55 | 0.1482 | -3.25 | 0.0015 |
| | 4 | -1.36 | 0.2024 | -2.75 | 0.0071 |

Notes: The log money-output ratio is computed by dividing the indicated monetary aggregate by nominal GDP and taking the natural logarithm. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A12. Efficient Unit Root Test Results v_t^a : income velocity

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | ADF-GLS | | ADF-GLS | |
|----------------------|-----|---------|------------|------------|------------|
| | | Level | p -value | Difference | p -value |
| M1 | 1 | 0.23 | 0.5902 | -4.13 | 0.0001 |
| | 2 | 0.13 | 0.5792 | -3.74 | 0.0003 |
| | 3 | 0.08 | 0.5662 | -3.42 | 0.0008 |
| | 4 | 0.03 | 0.5612 | -3.37 | 0.0015 |
| M2 | 1 | -0.77 | 0.4453 | -3.03 | 0.0040 |
| | 2 | -0.78 | 0.4475 | -2.88 | 0.0053 |
| | 3 | -0.67 | 0.4859 | -2.41 | 0.0207 |
| | 4 | -0.74 | 0.4650 | -2.20 | 0.0323 |
| MZM | 1 | -1.16 | 0.2558 | -4.18 | 0.0001 |
| | 2 | -0.86 | 0.3596 | -3.99 | 0.0002 |
| | 3 | -0.77 | 0.3891 | -3.63 | 0.0004 |
| | 4 | -0.77 | 0.3888 | -3.39 | 0.0006 |
| M4 | 1 | -1.58 | 0.1434 | -3.86 | 0.0003 |
| | 2 | -1.53 | 0.1503 | -3.34 | 0.0014 |
| | 3 | -1.52 | 0.1572 | -3.22 | 0.0017 |
| | 4 | -1.37 | 0.2091 | -2.80 | 0.0063 |

Notes: Income velocity is computed by dividing nominal GDP by the indicated monetary aggregate. ADF-GLS is the modified Dickey-Fuller test statistic proposed by Elliot, Rothenberg, and Stock (1996). The p -values from testing the null hypothesis of a unit root in the level or first difference of the indicated series are bootstrapped following Benati (2015). q denotes the number of lags included in the modified ADF regression.

Table A13. Cointegrating Money Demand RelationshipsMeltzer Specification: $\ln(m_t^a) = \alpha_0 - \alpha_1 \ln(r_t - r_t^a)$ $\ln(m_t^a)$: log money-output ratio

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | α_0 | α_1 | Max Eigenvalue $r = 0$ p -value | | Max Eigenvalue $r = 1$ p -value | |
|----------------------|-----------------|------------|------------|--------------------------------------|--------|--------------------------------------|--------|
| M1 | 2 ⁺⁺ | -6.69 | 1.94 | 44.81 | 0.0001 | 3.09 | 0.6588 |
| | 3 | -7.06 | 1.98 | 39.05 | 0.0001 | 3.12 | 0.6658 |
| | 4 | -6.62 | 1.78 | 26.80 | 0.0014 | 2.51 | 0.7545 |
| | 5 | -6.33 | 1.66 | 21.04 | 0.0135 | 2.10 | 0.8330 |
| M2 | 2 ⁺⁺ | -4.93 | 1.55 | 44.20 | 0.0001 | 5.53 | 0.2561 |
| | 3 | -4.98 | 1.57 | 43.44 | 0.0001 | 5.13 | 0.3019 |
| | 4 | -5.00 | 1.58 | 33.50 | 0.0001 | 5.18 | 0.3038 |
| | 5 | -5.12 | 1.67 | 30.70 | 0.0002 | 4.13 | 0.4378 |
| MZM | 2 ⁺⁺ | -4.70 | 1.38 | 39.38 | 0.0001 | 5.95 | 0.2078 |
| | 3 | -4.69 | 1.37 | 38.78 | 0.0001 | 5.43 | 0.2591 |
| | 4 | -4.78 | 1.42 | 27.95 | 0.0007 | 5.57 | 0.2495 |
| | 5 | -4.78 | 1.43 | 26.61 | 0.0012 | 4.49 | 0.3829 |
| M4 | 2 ⁺⁺ | -10.38 | 5.99 | 24.13 | 0.0026 | 8.62 | 0.0678 |
| | 3 | -8.79 | 4.71 | 21.81 | 0.0055 | 8.54 | 0.0689 |
| | 4 | -35.20 | 24.83 | 18.03 | 0.0309 | 8.07 | 0.0888 |
| | 5 | -11.73 | 6.96 | 15.37 | 0.0792 | 6.81 | 0.1467 |

Notes: The log money-output ratio is computed by dividing the indicated monetary aggregate by nominal GDP and taking the natural logarithm. α_0 and α_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table A14. Cointegrating Money Demand RelationshipsCagan Specification: $\ln(m_t^a) = \delta_0 - \delta_1(r_t - r_t^a)$ $\ln(m_t^a)$: log money-output ratio

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | δ_0 | δ_1 | Max Eigenvalue $r = 0$ p -value | | Max Eigenvalue $r = 1$ p -value | |
|----------------------|-----|------------|------------|--------------------------------------|--------|--------------------------------------|--------|
| M1 | 2* | 0.88 | 28.02 | 30.17 | 0.0007 | 1.87 | 0.8501 |
| | 3+ | 2.29 | 40.01 | 28.11 | 0.0007 | 2.90 | 0.6723 |
| | 4 | 1.12 | 30.39 | 17.60 | 0.0377 | 2.04 | 0.8214 |
| | 5 | 1.66 | 34.49 | 16.51 | 0.0622 | 1.65 | 0.8855 |
| M2 | 2* | -1.32 | 5.38 | 41.71 | 0.0001 | 5.54 | 0.2521 |
| | 3+ | -1.16 | 5.98 | 42.46 | 0.0001 | 5.22 | 0.2893 |
| | 4 | -1.19 | 5.92 | 31.89 | 0.0005 | 5.28 | 0.2881 |
| | 5 | -0.98 | 6.64 | 30.55 | 0.0003 | 4.33 | 0.3973 |
| MZM | 2 | -1.52 | 4.69 | 35.13 | 0.0001 | 5.94 | 0.2099 |
| | 3** | -1.38 | 5.24 | 39.11 | 0.0001 | 5.49 | 0.2681 |
| | 4 | -1.40 | 5.22 | 26.22 | 0.0016 | 5.54 | 0.2523 |
| | 5 | -1.29 | 5.58 | 26.91 | 0.0015 | 4.53 | 0.3710 |
| M4 | 2** | 1.30 | 13.86 | 21.39 | 0.0063 | 8.66 | 0.0617 |
| | 3 | 0.37 | 10.74 | 21.09 | 0.0099 | 8.17 | 0.0819 |
| | 4 | 8.52 | 40.36 | 16.04 | 0.0628 | 8.01 | 0.0861 |
| | 5 | 2.15 | 17.11 | 15.30 | 0.0779 | 6.65 | 0.1458 |

Notes: The log money-output ratio is computed by dividing the indicated monetary aggregate by nominal GDP and taking the natural logarithm. δ_0 and δ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table A15. Cointegrating Money Demand RelationshipsSelden-Latané Specification: $v_t^a = \gamma_0 + \gamma_1(r_t - r_t^a)$ v_t^a : income velocity

Sample Period: 1967:1 – 2017:2

| Divisia Aggregate | q | γ_0 | γ_1 | Max Eigenvalue $r = 0$ p -value | | Max Eigenvalue $r = 1$ p -value | |
|----------------------|-----|------------|------------|--------------------------------------|--------|--------------------------------------|--------|
| M1 | 2* | -31.40 | 365.06 | 26.25 | 0.0016 | 2.20 | 0.7836 |
| | 3+ | -86.89 | 835.35 | 24.42 | 0.0031 | 3.52 | 0.5564 |
| | 4 | -40.37 | 446.39 | 15.50 | 0.0747 | 2.45 | 0.7453 |
| | 5 | -33.32 | 381.88 | 15.72 | 0.0734 | 2.01 | 0.8218 |
| M2 | 2* | -14.01 | 111.02 | 33.39 | 0.0001 | 3.90 | 0.4505 |
| | 3+ | -18.64 | 128.67 | 33.68 | 0.0001 | 3.83 | 0.4671 |
| | 4 | -19.44 | 132.49 | 25.73 | 0.0021 | 3.76 | 0.4901 |
| | 5 | -20.11 | 133.89 | 26.28 | 0.0016 | 3.24 | 0.5664 |
| MZM | 2 | -10.33 | 98.68 | 28.68 | 0.0008 | 3.97 | 0.4202 |
| | 3** | -15.17 | 117.60 | 31.46 | 0.0002 | 4.10 | 0.4376 |
| | 4 | -17.47 | 127.05 | 20.90 | 0.0105 | 4.02 | 0.4328 |
| | 5 | -14.41 | 114.42 | 22.81 | 0.0062 | 3.46 | 0.5291 |
| M4 | 2** | -1678.65 | 5975.83 | 20.53 | 0.0099 | 6.67 | 0.1512 |
| | 3 | -354.33 | 1322.00 | 19.42 | 0.0163 | 6.47 | 0.1677 |
| | 4 | 61.53 | -173.20 | 15.21 | 0.0822 | 6.13 | 0.1900 |
| | 5 | 371.27 | -1279.26 | 14.32 | 0.1134 | 5.61 | 0.2322 |

Notes: Income velocity is computed by dividing nominal GDP by the indicated monetary aggregate. γ_0 and γ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on v_t^a equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table A16. Cointegrating Money Demand RelationshipsMeltzer Specification: $\ln(m_t^a) = \alpha_0 - \alpha_1 \ln(r_t - r_t^a)$ $\ln(m_t^a)$: log money-output ratio

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | α_0 | α_1 | Max Eigenvalue $r = 0$ p -value | | Max Eigenvalue $r = 1$ p -value | |
|----------------------|-----------------|------------|------------|--------------------------------------|--------|--------------------------------------|--------|
| M1 | 2 ⁺⁺ | -5.31 | 1.17 | 33.31 | 0.0002 | 2.11 | 0.8362 |
| | 3 | -6.50 | 1.70 | 25.24 | 0.0031 | 1.82 | 0.8845 |
| | 4 | -11.64 | 3.97 | 20.58 | 0.0145 | 1.85 | 0.8885 |
| | 5 | -17.52 | 6.60 | 13.66 | 0.1578 | 1.93 | 0.8795 |
| M2 | 2 ⁺⁺ | -3.77 | 0.70 | 27.33 | 0.0009 | 0.65 | 0.9917 |
| | 3 | -3.81 | 0.72 | 21.27 | 0.0110 | 0.73 | 0.9892 |
| | 4 | -3.89 | 0.78 | 18.13 | 0.0331 | 1.14 | 0.9459 |
| | 5 | -4.02 | 0.88 | 14.90 | 0.1015 | 1.57 | 0.8848 |
| MZM | 2 ⁺⁺ | -3.92 | 0.79 | 28.76 | 0.0002 | 0.94 | 0.9680 |
| | 3 | -3.95 | 0.81 | 17.88 | 0.0331 | 1.42 | 0.8988 |
| | 4 | -3.98 | 0.82 | 14.68 | 0.1029 | 2.11 | 0.7781 |
| | 5 | -3.96 | 0.81 | 12.00 | 0.2326 | 3.57 | 0.5006 |
| M4 | 2 ⁺⁺ | -11.40 | 6.69 | 16.89 | 0.0472 | 4.90 | 0.3059 |
| | 3 | 5.02 | -5.79 | 13.70 | 0.1302 | 4.29 | 0.3766 |
| | 4 | 0.72 | -2.52 | 10.71 | 0.3255 | 3.92 | 0.4309 |
| | 5 | 1.80 | -3.36 | 8.97 | 0.4927 | 3.59 | 0.4905 |

Notes: The log money-output ratio is computed by dividing the indicated monetary aggregate by nominal GDP and taking the natural logarithm. α_0 and α_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table A17. Cointegrating Money Demand RelationshipsCagan Specification: $\ln(m_t^a) = \delta_0 - \delta_1(r_t - r_t^a)$ $\ln(m_t^a)$: log money-output ratio

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | δ_0 | δ_1 | Max Eigenvalue $r = 0$ p -value | | Max Eigenvalue $r = 1$ p -value | |
|----------------------|-----------------|------------|------------|--------------------------------------|--------|--------------------------------------|--------|
| M1 | 2 ⁺⁺ | 0.21 | 25.26 | 32.17 | 0.0001 | 1.67 | 0.8969 |
| | 3 | 3.15 | 51.72 | 27.28 | 0.0015 | 2.28 | 0.8082 |
| | 4 | -9.86 | -64.23 | 22.27 | 0.0067 | 2.29 | 0.8081 |
| | 5 | -6.32 | -32.60 | 16.89 | 0.0576 | 3.03 | 0.6884 |
| M2 | 2 ⁺⁺ | -1.90 | 3.38 | 24.92 | 0.0027 | 0.62 | 0.9926 |
| | 3 | -1.92 | 3.31 | 19.32 | 0.0218 | 0.77 | 0.9811 |
| | 4 | -1.83 | 3.61 | 17.49 | 0.0399 | 1.42 | 0.8995 |
| | 5 | -1.86 | 3.51 | 16.07 | 0.0721 | 2.09 | 0.7839 |
| MZM | 2 ⁺⁺ | -1.83 | 3.81 | 26.48 | 0.0014 | 0.94 | 0.9703 |
| | 3 | -1.81 | 3.87 | 17.88 | 0.0303 | 1.42 | 0.8977 |
| | 4 | -1.77 | 4.01 | 15.61 | 0.0775 | 2.26 | 0.7420 |
| | 5 | -1.90 | 3.58 | 13.64 | 0.1556 | 3.90 | 0.4392 |
| M4 | 2 ⁺⁺ | -21.71 | -68.87 | 16.35 | 0.0500 | 5.02 | 0.2871 |
| | 3 | -5.95 | -12.10 | 13.73 | 0.1310 | 4.32 | 0.3828 |
| | 4 | -4.29 | -6.08 | 10.57 | 0.3302 | 3.94 | 0.4279 |
| | 5 | -4.84 | -8.01 | 9.91 | 0.3959 | 3.42 | 0.5171 |

Notes: The log money-output ratio is computed by dividing the indicated monetary aggregate by nominal GDP and taking the natural logarithm. δ_0 and δ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on $\ln(m_t^a)$ equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

Table A18. Cointegrating Money Demand RelationshipsSelden-Latané Specification: $v_t^a = \gamma_0 + \gamma_1(r_t - r_t^a)$ v_t^a : income velocity

Sample Period: 1983:1 – 2017:2

| Divisia Aggregate | q | γ_0 | γ_1 | Max Eigenvalue | | Max Eigenvalue | |
|----------------------|-----------------|------------|------------|----------------|------------|----------------|------------|
| | | | | $r = 0$ | p -value | $r = 1$ | p -value |
| M1 | 2 ⁺⁺ | -0.24 | 125.92 | 30.00 | 0.0005 | 1.88 | 0.8570 |
| | 3 | -2.16 | 144.55 | 23.46 | 0.0043 | 1.88 | 0.8559 |
| | 4 | -9.62 | 209.91 | 19.09 | 0.0275 | 1.74 | 0.8738 |
| | 5 | -14.99 | 256.83 | 15.04 | 0.1024 | 2.22 | 0.8114 |
| M2 | 2 ⁺⁺ | 4.02 | 47.26 | 25.06 | 0.0027 | 0.57 | 0.9951 |
| | 3 | 4.31 | 46.39 | 19.22 | 0.0222 | 0.56 | 0.9957 |
| | 4 | 3.49 | 49.25 | 17.23 | 0.0467 | 0.83 | 0.9759 |
| | 5 | 3.67 | 48.74 | 15.67 | 0.0819 | 1.21 | 0.9310 |
| MZM | 2 ⁺⁺ | 2.40 | 55.88 | 26.20 | 0.0020 | 0.88 | 0.9708 |
| | 3 | 2.35 | 56.50 | 17.79 | 0.0355 | 1.04 | 0.9564 |
| | 4 | 1.88 | 58.23 | 15.10 | 0.0936 | 1.50 | 0.8892 |
| | 5 | 3.27 | 53.37 | 12.96 | 0.1799 | 2.31 | 0.7296 |
| M4 | 2 ⁺⁺ | -165.74 | 645.98 | 16.50 | 0.0483 | 4.54 | 0.3363 |
| | 3 | 74.67 | -220.70 | 14.00 | 0.1215 | 4.07 | 0.4105 |
| | 4 | 41.13 | -99.68 | 10.68 | 0.3194 | 3.70 | 0.4700 |
| | 5 | 44.73 | -111.66 | 9.88 | 0.3996 | 3.30 | 0.5354 |

Notes: Income velocity is computed by dividing nominal GDP by the indicated monetary aggregate. γ_0 and γ_1 are Johansen's (1991) maximum likelihood estimates of the coefficients of the first cointegrating vector, normalized so that the coefficient on v_t^a equals one. The maximum eigenvalue test statistics for the null hypothesis of $r = 0$ and $r = 1$ cointegrating vectors are shown, together with bootstrapped p -values, computed using the algorithm outlined by Cavaliere, Rahbek, and Taylor (2012). q denotes the number of lags included in the VAR in levels; the superscripts * and + denote the optimal lag lengths chosen according to the Schwarz (1978) and Hannan-Quinn (1979) criteria.

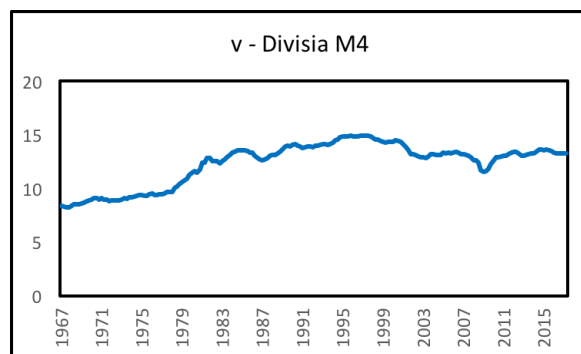
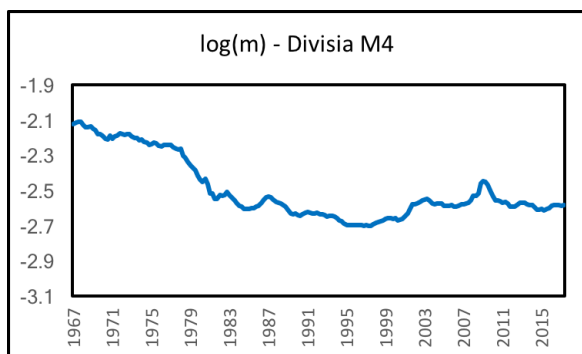
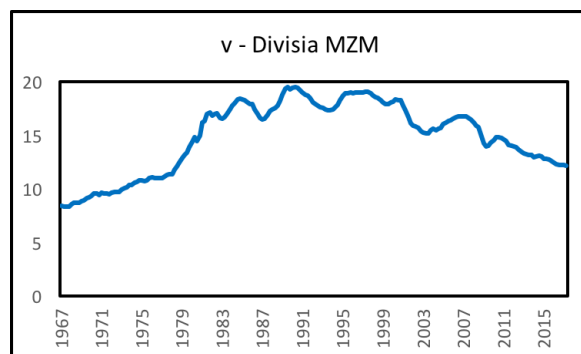
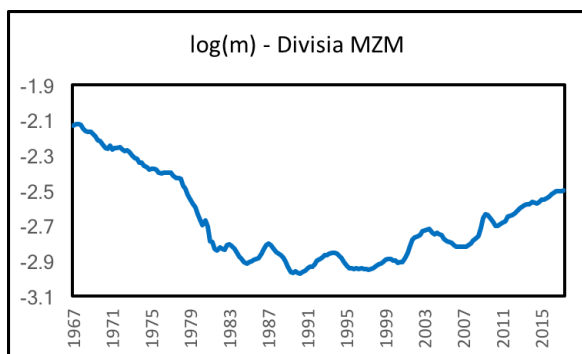
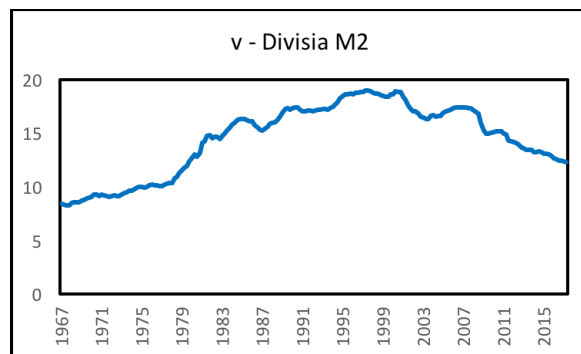
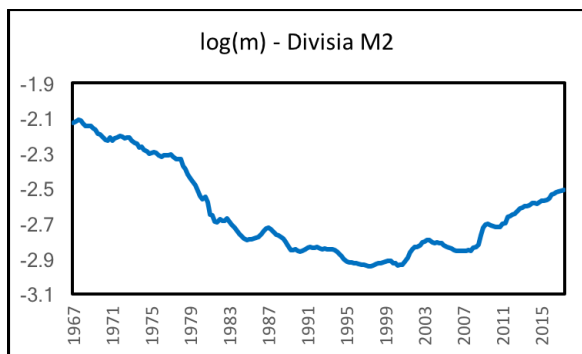
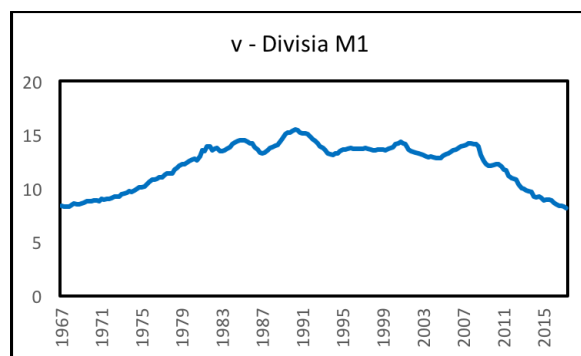
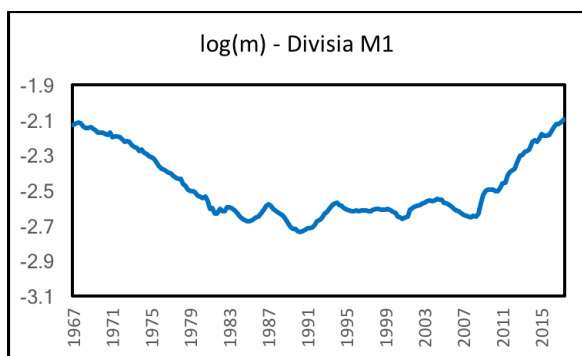


Figure A1. Divisia Monetary Data. The log money-output ratio in column one and income velocity in column two are computed using nominal GDP as the scale variable.

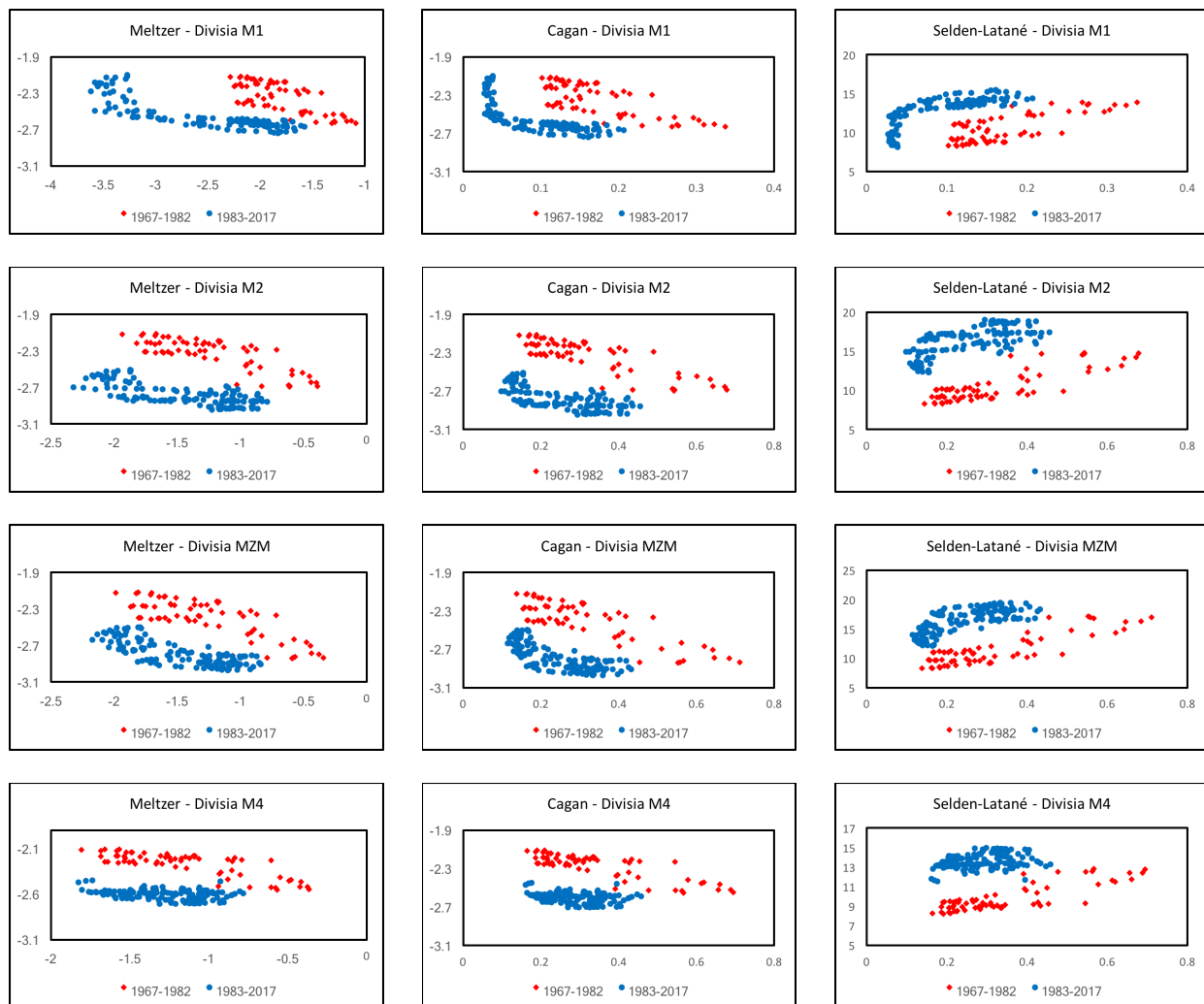


Figure A2. Divisia Money Demand Relationships. Each scatterplot compares the indicated Divisia monetary aggregate, scaled by nominal GDP, measured along the vertical axis, to the associated user cost aggregate, measured along the horizontal axis. The Meltzer (1963) specification relates the log money-output ratio to the log of the user cost; the Cagan (1956) specification relates the log money-output ratio to the level of the user cost; the Selden-Latané specification relates the income velocity of money to the level of the user cost.